



# Fixed-dose combination antituberculosis therapy: a systematic review and meta-analysis

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**ABSTRACT** Fixed-dose combination (FDC) formulations are currently recommended for the treatment of active tuberculosis (TB). We have conducted a systematic review to evaluate the risk of treatment failure or disease relapse, acquired drug resistance, bacterial conversion after 2 months of treatment, adverse events, adherence and treatment satisfaction associated with treatment of active TB using FDC or separate drug formulations.

We searched four electronic databases for randomised controlled trials and cohort studies. Results from trials that directly compared FDC to separate drug formulations were pooled. Results from other studies were reported separately.

We identified 2450 citations from which 15 controlled trials and four additional relevant studies were included. In the 15 trials there were no differences in acquired drug resistance, bacterial conversion after 2 months of treatment or adverse drug reactions with FDC or separate drug formulations. There was a trend toward higher risk of failure or relapse with FDC (pooled relative risk 1.28 (95% CI 0.99–1.7)). Based on individual study results, only one of two trials that assessed treatment satisfaction, and none of five that assessed patient adherence, favoured FDCs.

Although FDC formulations simplify TB therapy, the current evidence does not indicate that these formulations improve treatment outcomes among patients with active TB.



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Current evidence does not indicate that fixed-dose combinations of first-line TB drugs improve treatment outcomes <http://ow.ly/la48v>

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This article has supplementary material available from [www.erj.ersjournals.com](http://www.erj.ersjournals.com)

Received: Nov 08 2012 | Accepted after revision: Dec 20 2012 | First published online: Jan 11 2013

Support statement: A.S. Albanna received salary support from the Ministry of Higher Education, Government of Saudi Arabia.

Conflict of interest: None declared.

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## Introduction

Tuberculosis (TB) is a global health problem, with 8.7 million new cases and accounting for ~1.4 million deaths annually [1]. Moreover, strains of *Mycobacterium tuberculosis* that are resistant to standard anti-TB therapy are emerging in almost all areas reporting to the World Health Organization (WHO) [2]. Nonadherence to treatment regimen and inappropriate prescription of TB therapy are believed to be major contributing factors to this public health problem [3, 4]. Due to the large number of tablets used in the treatment regimens of TB, fixed-dose combination (FDC) tablets, each combining two or more anti-TB drugs, have been manufactured since the 1980s [5] to simplify TB therapy and facilitate physician and patient compliance with treatment recommendations [6]. These FDC tablets also prevent inadvertent monotherapy, which may occur because of physician error in prescription, inadequate regimens or patient error in selectively taking only one drug. In addition, dealing with one combined formulation that contains all essential drugs simplifies drug procurement, storage and distribution, and may consequently reduce drug supply management errors and cost.

In 1994, the WHO and the International Union against Tuberculosis and Lung Disease (IUATLD) recommended the use of FDC anti-TB therapy [7]. Following the announcement of this recommendation, and its more widespread implementation, concerns were raised about adequate bioavailability of the component drugs, particularly rifampicin (RIF) due to its enhanced decomposition in the presence of isoniazid (INH) [8–11]. As a result, the WHO and the IUATLD established guidelines for assuring the bioavailability of FDC anti-TB drug components [12]. Currently, the WHO Model List of Essential Drugs includes two-drug formulations (INH + RIF and INH + ethambutol), three-drug formulations (INH + RIF + ethambutol and INH + RIF + pyrazinamide) and a four-drug formulation (INH + RIF + ethambutol + pyrazinamide) [13].

Despite the anticipated advantages of FDC anti-TB drugs, questions about their effectiveness have not been answered. Many observational studies and clinical trials have been conducted to assess the effectiveness of FDC drugs in reducing treatment failure, disease relapse and the emergence of drug resistance. Among these studies, the use of FDC drugs has resulted in favourable [14], unfavourable [15] or unchanged treatment outcomes [16, 17].

Due to the anticipated advantages, and despite the current conflicting evidence, the FDC formulations are recommended for treatment of active TB by the WHO [18], the International Standards for TB Care (Standard 8) [19], and the American Thoracic Society [20].

## Study questions

The study aimed to answer the following: 1) in patients who are treated for bacteriologically confirmed TB, is anti-TB therapy using FDC drug formulations associated with lower rates of bacteriologically confirmed treatment failure, disease relapse or emergence of drug resistance when compared to separate-drug formulations?; and 2) in patients receiving TB treatment, are adverse drug reactions, patient adherence and treatment satisfaction superior with FDC than separate-drug formulations?

## Methods

### Search strategy and study selection

A search strategy was designed to retrieve articles investigating FDC anti-TB therapy published in any language between January 1980 and July 2011. The databases used for the literature search were MEDLINE (Ovid platform); MEDLINE In-Process & Other Non-Indexed Citations (Ovid platform); Embase (Ovid platform); The Cochrane Library (published by Wiley), which includes Cochrane Reviews, DARE and Central Register of Controlled Clinical Trials; and LILACS (BIREME, PAHO and WHO Latin-American and Caribbean Center on Health Sciences Information) databases. The following four sets of search terms were combined with “AND”: 1) terms about TB, *Mycobacterium* and anti-TB; 2) terms to restrict for treatment regimens that contain both isoniazid and rifampicin; 3) terms to restrict for the use of combination formulations; and 4) restriction to human studies published since 1980. For more details about the terms used in each database, refer to the online supplementary material.

Studies that fulfilled all of the following criteria were eligible for full-text review: 1) randomised clinical trial (RCT) or cohort study (the latter should include  $\geq 50$  subjects); 2) bacteriologically confirmed diagnosis of active TB, based on culture or smear analyses, among included subjects; 3) treatment with an FDC anti-TB formulation that contained at least RIF and INH; 4) treatment with an effective anti-TB regimen (*i.e.* daily or at least three times weekly administration of RIF and INH for 9 months, or for 6 months when pyrazinamide was added during the initial 2 months); 5) measurement of at least one of our primary treatment outcomes (*i.e.* bacteriologically confirmed treatment failure or relapse, or acquired drug

resistance with diagnosis based on baseline and follow-up drug sensitivity testing); and 6) follow-up period of  $\geq 5$  months during the treatment.

Selection of eligible studies was performed in a stepwise fashion: titles, then abstracts, then full texts, by two reviewers (A. Albanna and B. Smith) working independently. At each stage, all studies selected by either reviewer (*i.e.* concordant eligible or discordant) were included for full-text review. Inclusion of studies, after full-text review, was based on concordance of the two reviewers; disagreement was resolved by a third reviewer (D. Menzies).

#### **Data extraction**

The extracted data included information about the context of the study (study design, location and time period), characteristics of included subjects (age, sex, past TB treatment, HIV status and comorbidities), disease status (disease site and drug sensitivity) and treatment outcomes (completion of treatment, compliance to treatment, adverse drug reaction, treatment failure, death during treatment, disease relapse, acquired drug resistance and patient satisfaction). In addition, a quality assessment scale was adapted from the Cochrane Collaboration tool to assess the following five quality indicators: 1) sequential or randomised allocation of subjects to study groups; 2) concealment of the allocation, in case of RCTs; 3) adequate assessment of incomplete outcome data; 4) reporting of pre-specified or all expected outcomes (to obviate the possibility of selective outcome reporting); and 5) adequate consideration of potential sources of bias. To ensure accurate and consistent data collection, both reviewers independently performed data extraction from a sample of nine articles. Important missing data were obtained by correspondence with the studies' authors through email contact.

#### **Outcome measures**

The pre-specified primary outcome measures were "treatment failure or disease relapse", as one outcome, and acquired drug resistance as another. The pre-specified secondary outcomes were bacterial conversion after 2 months of treatment, adverse drug reaction, patient adherence and treatment satisfaction. Pre-specified subgroup analysis was stratified by baseline drug sensitivity testing, study quality, publication year, treatment supervision modality, type of treatment regimen and FDC formulation/producer. Our decision to stratify the studies by their potential conflict of interest was made after collecting the data (*post hoc* analysis).

#### **Data analysis**

Differences in the outcomes between the comparative groups were expressed as risk ratios and 95% confidence intervals, using per-protocol analysis. The effect measures of comparative RCTs were pooled using the DerSimonian–Laird random effects model. The use of a random effects, rather than a fixed effect, model was pre-specified to account for variations between studies related to the type and severity of prevalent disease, standard of care and research quality. To obtain valid, unbiased comparative estimates, our analysis focused on the comparative RCTs, which represented the majority of the included studies. Summaries of the effect measures from the other studies were not pooled and were reported separately. Between-study heterogeneity was assessed using Chi-squared (Cochran's Q), indicating statistical significance as  $p < 0.1$  and I-squared tests. The latter are interpreted as showing unimportant heterogeneity if values are  $< 40\%$ , moderate heterogeneity if values are between 40% and 60% and substantial heterogeneity if values  $> 60\%$ . In the case of moderate or substantial heterogeneity of results, or inconsistent methods of ascertainment across studies, the outcome estimates were not pooled and were reported separately. Subgroup and meta-regression analyses were performed to detect factors that influenced the primary outcome results. Reporting bias, which includes publication bias, was assessed using funnel plot and Egger's test, which is based on linear regression analysis to test the association between the intervention effect (using logarithmic scale) and its standard error [21]. All analyses were conducted using STATA (version 12) (StataCorp, College Station, TX, USA) software.

#### **Results**

Of 2450 citations identified by our search strategy, 25 met the inclusion criteria for this review. These 25 articles reported results of 19 different studies (fig. 1). Among these 19 studies, 15 RCTs directly compared FDC to separate drug formulations and included a total of 5630 subjects (table 1). The other four studies represent one comparative cohort [39], two noncomparative (*i.e.* no direct comparison between FDC and separate drug formulations) RCTs [40–42] and one noncomparative cohort [43] which included total numbers of 474, 310 and 1888 subjects, respectively; refer to the online supplementary material for study descriptions.

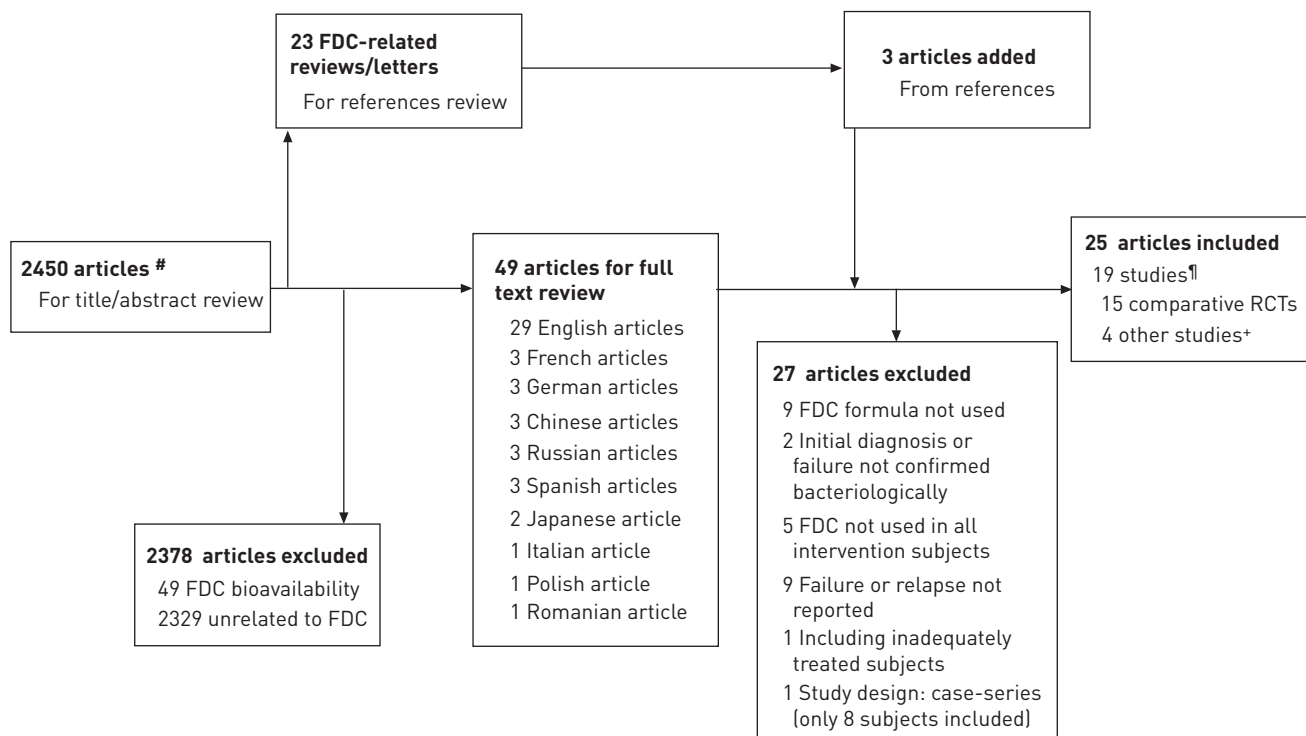


FIGURE 1 Study selection. FDC: fixed-dose combination; RCTs: randomised controlled trials. #: After excluding duplicate articles; ¶: some studies were published in more than one article; +: one comparative cohort and three noncomparative studies.

### Primary outcome results of the comparative RCTs

In the 15 RCTs there was a trend toward higher risk of treatment failure or disease relapse with FDC compared to separate drug formulations (pooled relative risk 1.28 (95% CI 0.99–1.7)), with no significant heterogeneity between the results of different studies (fig. 2). The incidence of failure or relapse was relatively low in both treatment arms (table 2), and the pooled risk difference was 1% (95% CI -0.2–2%) higher with FDCs.

As seen in table 2, the risk of acquired drug resistance, based on pooled results from four RCTs, was very low in both treatment arms and the relative risk estimate was inconclusive.

In the subgroup analyses, baseline drug sensitivity status appeared to modify the risk of “treatment failure or disease relapse”. Comparing FDC with separate-drug formulations, the risk was significantly higher with FDCs within the stratum of subjects with baseline drug-susceptible TB (pooled risk ratio 1.48 (95% CI 1.04–2)), and lower, although not significantly so, with FDCs within the drug-resistant stratum. In addition, FDC formulation was inferior to separate-drug formulation among subjects receiving self-administered therapy and in studies with no potential conflict of interest (fig. 3).

Univariate metaregression analyses did not indicate a significant influence of publication year or study quality on the outcome results (fig. 4). After including these two covariates with drug susceptibility, treatment supervision and potential conflict of interest variables in a multivariate metaregression model, drug susceptibility was the only variable that significantly modified the outcome results (comparing the point estimate within drug-resistant to the point estimate within drug-susceptible strata, risk ratio 0.32 (95% CI 0.11–0.94);  $p=0.04$ ).

Funnel plot analysis demonstrated a symmetrical distribution of “treatment failure or relapse” effect estimates across studies and the regression line indicated that small studies, which have less precise estimates (larger standard errors), tended to shift the treatment effect in favour of FDC treatment (fig. 5). However, the small-study effect was not significant (estimated bias coefficient -0.36 (95% CI -1.2–0.49);  $p=0.39$ ).

### Secondary outcome results of comparative RCTs

As seen in table 2, FDC treatment was almost similar to separate-formulation treatment for eliminating mycobacterial isolation after 2 months of treatment and had similar association with adverse drug reaction. The estimated results of patient adherence and treatment satisfaction outcomes were not pooled because of

TABLE 1 Summary of the comparative randomised controlled trials studied

Author [Ref.]	Year of publication	Location of study	Mean age years	Male %	Treatment regimen	FDC formulation	DOT	Proper allocation sequence	Allocation concealment	Follow-up completion <sup>#</sup>	Non-selective outcomes <sup>†</sup>	Free of bias <sup>*</sup>
<b>RCTAI [22]</b>	1989	India	29 <sup>§</sup>	70	HRZ	Rifater/Rifinah	No	Yes	Unclear	Yes	Yes	Yes
<b>COMIE and BRINK [23]</b>	1990	South Africa	38	100	HRZ <sub>±</sub> S	Rifater	Yes	No <sup>†</sup>	Yes	Yes	Yes	No <sup>##</sup>
<b>HKCS and BMRC [24, 25]</b>	1991, 1989	China	35 <sup>§</sup>	66	HRZ <sub>±</sub> S	Rifater	Yes	Unclear	Unclear	Yes	Yes	Yes
<b>GLATTHAAR et al. [26]</b>	1991	South Africa	NS	NS	HRZE	Rifater	Yes	Unclear	Unclear	Yes	Yes	Unclear
<b>MACNAB et al. [27]</b>	1994	South Africa	NS	NS	HRZE	Rifater	Yes	No <sup>†</sup>	Unclear	No <sup>†*</sup>	Yes	Unclear
<b>CHAULET and colleagues [28-30]</b>	1995, 1990, 1989	Algeria	28 <sup>§</sup>	75	HRZ <sup>++</sup>	NS	No <sup>§§</sup>	Unclear	Unclear	Yes	Yes	Yes
<b>ZHANG et al. [31]</b>	1996	China	41 <sup>§</sup>	65	HRZ	Rifater/Rifinah	Yes	Yes	Unclear	Yes	Yes	Yes
<b>ZHU et al. [32]</b>	1998	China	37 <sup>§</sup>	70	HRZ	Rifater/Rifinah	NS	Unclear	Unclear	Yes	Yes	Yes
<b>Teo [33] and STS and BMRC [34]</b>	1999, 1991	Singapore	39 <sup>§</sup>	66	HRZ <sub>±</sub> S	Rifater	Yes	Unclear	Unclear	Yes	Yes	Yes
<b>SU and PERNÉ [35]</b>	2002	Taiwan	NS	89	HRZ	Rifater/Rifinah	No	Unclear	Unclear	No <sup>†*</sup>	Yes	Unclear
<b>MUNTEANU et al. [36]</b>	2004	Romania	37 <sup>§</sup>	63	HRZE	NS	Yes <sup>†/</sup>	Unclear	Unclear	Yes	Yes	Yes
<b>XU et al. [37]</b>	2004	China	49	76	HRZE	NS	NS	No <sup>†</sup>	Unclear	Yes	Yes	Unclear
<b>SURYANTO et al. [15] and GRAVENEEL et al. [38]</b>	2008, 2003	Indonesia	37	57	HRZE	Swizera	No	Yes	No <sup>###</sup>	Yes	Yes	Unclear
<b>BARTACEK et al. [17]</b>	2009	5 countries <sup>***</sup>	37	69	HRZE	Rimstar/ Rimactazid	NS	Yes	Yes	Yes	Yes	Yes
<b>LIENHARDT et al. [16]</b>	2011	9 countries <sup>+++</sup>	34	67	HRZE <sup>++</sup>	Swizera	Yes	Yes	Yes	Yes	Yes	Yes

FDC: fixed dose combination; DOT: direct observed therapy; RCTAI: Research Committee of the Tuberculosis Association of India; HKCS: Hong Kong Chest Service; BMRC: British Medical Research Council; STS: Singapore Tuberculosis Service; H: isoniazid; R: rifampicin; Z: pyrazinamide; S: streptomycin; E: ethambutol; NS: not specified. <sup>#</sup>: Complete follow-up for ≥75% of subjects, and assessment of the reasons for incomplete follow-up; <sup>†</sup>: free of selective outcome [i.e. reporting all expected or pre-specified outcomes]; <sup>†\*</sup>: equivalent subject characteristics and management between comparison groups, and the sample population had no specific risks that could influence their treatment outcomes; <sup>†/</sup>: the mean was estimated from a stratified age distribution; <sup>†/</sup>: allocation based on even versus odd generated numbers; <sup>##</sup>: streptomycin was added to the treatment of only one of the two groups; <sup>\*\*</sup>: <75% of subjects completed the follow-up; <sup>++</sup>: during the continuation phase, FDC was given to both groups; <sup>§</sup>: treatments were under direct supervision only during the first 3 weeks of therapy; <sup>†/</sup>: DOT was given only during the initial phase of treatment; <sup>###</sup>: the subjects were alternatively allocated to each study group; <sup>\*\*\*</sup>: Egypt, India, Pakistan, Philippines and Thailand; <sup>+++</sup>: Algeria, Colombia, Guinea, Vietnam, Nepal, Peru, Mozambique, Tanzania and Bolivia.

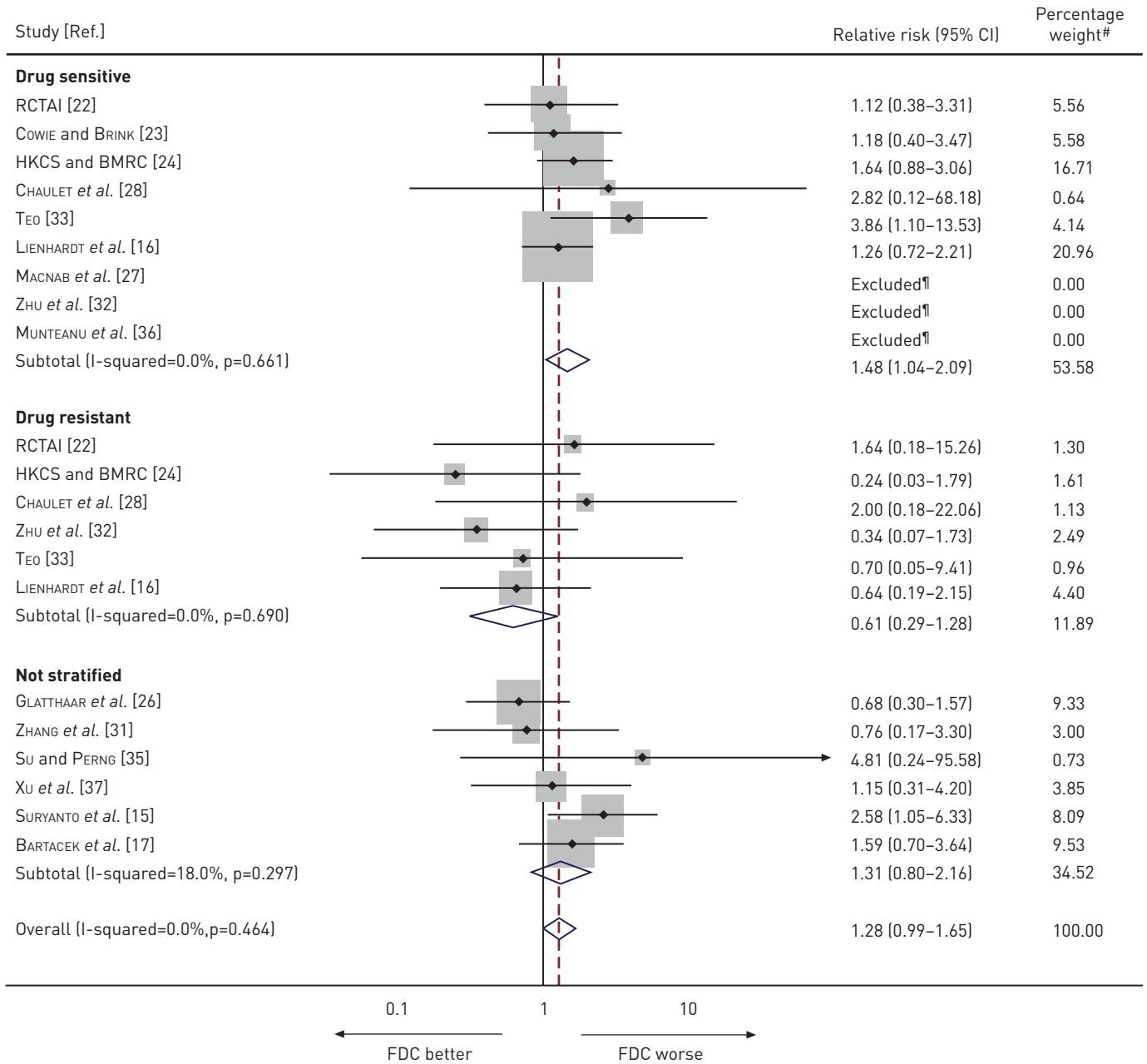


FIGURE 2 Forest plot of risk ratios of failure or relapse (main outcome) among fixed-dose combination (FDC) versus separate drug formulation groups, stratified by baseline drug susceptibility testing. RCTAI: Research Committee of the Tuberculosis Association of India; HKCS: Hong Kong Chest Service; BMRC: British Medical Research Council. <sup>#</sup>: From random effects analysis. <sup>††</sup>: Zero events in both arms, hence risk ratio was not estimated. When including these studies and adding 0.5 to each cell of the 2 × 2 table, the pooled risk ratio of the randomised controlled trials within the drug-sensitive stratum was 1.45 (95% CI 1.03–2.04), and the overall risk ratio was 1.26 (95% CI 0.98–1.63).

inconsistent ascertainment methods and significant heterogeneity of results (I-squared 67% and 98%, respectively) across the included RCTs. Only one of two RCTs that assessed treatment satisfaction, and none of five that assessed patient adherence, favoured FDCs.

**Outcome results of the cohort and noncomparative studies**

Among included studies, the comparative cohort [39] presented the highest proportion of “treatment failure or disease relapse” outcome, ranging from 5% to 11% among drug-susceptible and from 21% to 35% among drug-resistant TB patients. The crude risk ratio comparing FDC to separate-formulation treatments was 0.46 (95% CI 0.2–0.98) among drug-susceptible and 0.6 (95% CI 0.2–1.5) among drug-resistant TB patients. Results from the noncomparative studies [40–43] indicated a low proportion of

TABLE 2 Outcome results of comparative randomised controlled trials (RCTs)

Outcomes	Studies	FDC		Separate drug formulation		Risk ratio (95% CI)	Heterogeneity	
		Subjects	% (95% CI)	Subjects	% (95% CI)		I-squared	p-value
<b>Comparative RCTs (pooled)</b>								
Treatment failure or disease relapse	15	2750	4.2 (2.6–5.8)	2880	3.1 (1.9–4.2)	1.28 (0.99–1.7)	0	0.46
Acquired drug resistance	4	1113	0.26 (0–0.7)	1405	0.08 (0–0.35)	1.6 (0.5–5.4)	0	0.04
Tuberculosis culture conversion after 2 months of treatment	12	2354	94 (91–96)	2443	91 (89–92)	1.03 (1.01–1.04)	13	0.32
Adverse drug reaction	10	2416	16 (9–23)	2195	20 (11–28)	0.88 (0.75–1.03)	23.7	0.23
<b>Patients' adherence to treatment<sup>#</sup></b>								
RCTAI [22] <sup>†</sup>	1	95	77 (67–85)	101	73 (64–82)	1.05 (0.89–1.23)	66.5	0.02
COWIE and BRINK [23] <sup>†</sup>	1	69	58 (46–70)	81	84 (74–91)	0.69 (0.55–0.86)		
MACNAB et al. [27] <sup>§</sup>	1	121	65 (55–73)	79	57 (45–68)	1.13 (0.90–1.43)		
TEO [33]	1	154	95 (90–98)	153	97 (93–99)	0.97 (0.93–1.02)		
SU and PERNG [35] <sup>‡</sup>	1	57	70 (57–82)	48	67 (52–80)	1.05 (0.81–1.37)		
<b>Treatment satisfaction<sup>#</sup></b>								
TEO [33] <sup>##</sup>	1	154	92 (86–95)	153	90 (84–94)	1.02 (0.95–1.09)	97.8	0.00
BARTACEK et al. [17] <sup>¶¶</sup>	1	411	81 (77–85)	422	57 (52–61)	1.43 (1.30–1.58)		

Data are presented as n, unless otherwise stated. FDC: fixed-dose combination; RCTAI: Research Committee of the Tuberculosis Association of India. <sup>#</sup>: Results were not pooled because of significant heterogeneity between them and because of inconsistent methods for measurement of the outcome; <sup>†</sup>: assessment of adherence was based on monthly home visits and count of the number of remaining capsules; <sup>‡</sup>: assessment of adherence was based on urine tests and reports from medical staff; <sup>§</sup>: assessment of adherence was based on completion of  $\geq 75\%$  of the treatment doses; <sup>||</sup>: assessment of adherence was based on the loss of follow-up and alteration of treatment regimen; <sup>¶</sup>: assessment of satisfaction was based on spontaneous complaints; <sup>¶¶</sup>: assessment of satisfaction was based on patient's acceptance of the tablet number and size and complaints of swallowing problems.

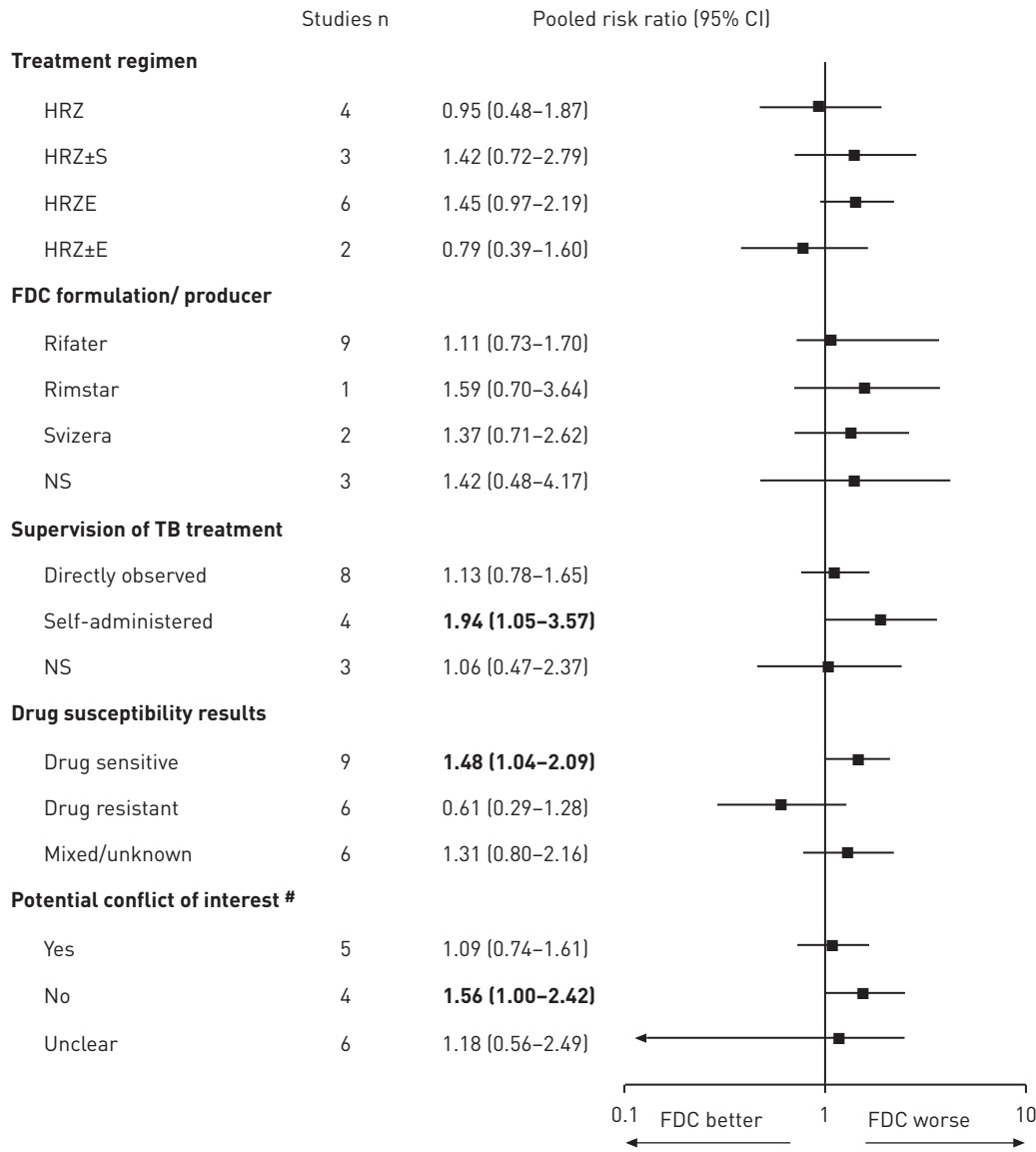


FIGURE 3 Subgroup analysis of the risk ratio of “treatment failure or disease relapse” among patients treated with fixed-dose combinations (FDC) or separate drug formulations. Bold type represents statistical significance. H: isoniazid; R: rifampicin; Z: pyrazinamide; S: streptomycin; E: ethambutol; TB: tuberculosis; NS: not specified. #: including funds and/or drug supplies.

“treatment failure or disease relapse”, ranging from 0.5% to 2%, and acquired drug resistance, ranging from 0 to 0.3%, among TB treated patients; for details refer to the online supplementary materials.

**Discussion**

Based on pooled results of RCTs, FDC therapy was associated with a trend toward increased risk of treatment failure or disease relapse, statistically insignificant difference in the emergence of drug resistance and adverse drug reactions and clinically unimportant difference in culture conversion after 2 months of treatment. Although one study identified better treatment satisfaction, none of the included studies identified better patient adherence among TB patients treated with FDC compared to separate drug formulations.

While the pooled result of the RCTs suggests that FDC treatment does not reduce the risk of failure or relapse (risk ratio estimate with a lower 95% CI range of 0.99 (close to the null value of 1.0)), it suggests potential increase in this risk (risk ratio estimate with an upper 95% CI range of 1.7). This could be explained by reduced bioavailability of FDC component drugs [8–12], when compared to separate-drug formulations. Because these outcomes were infrequent, the absolute increased risk of failure or relapse with



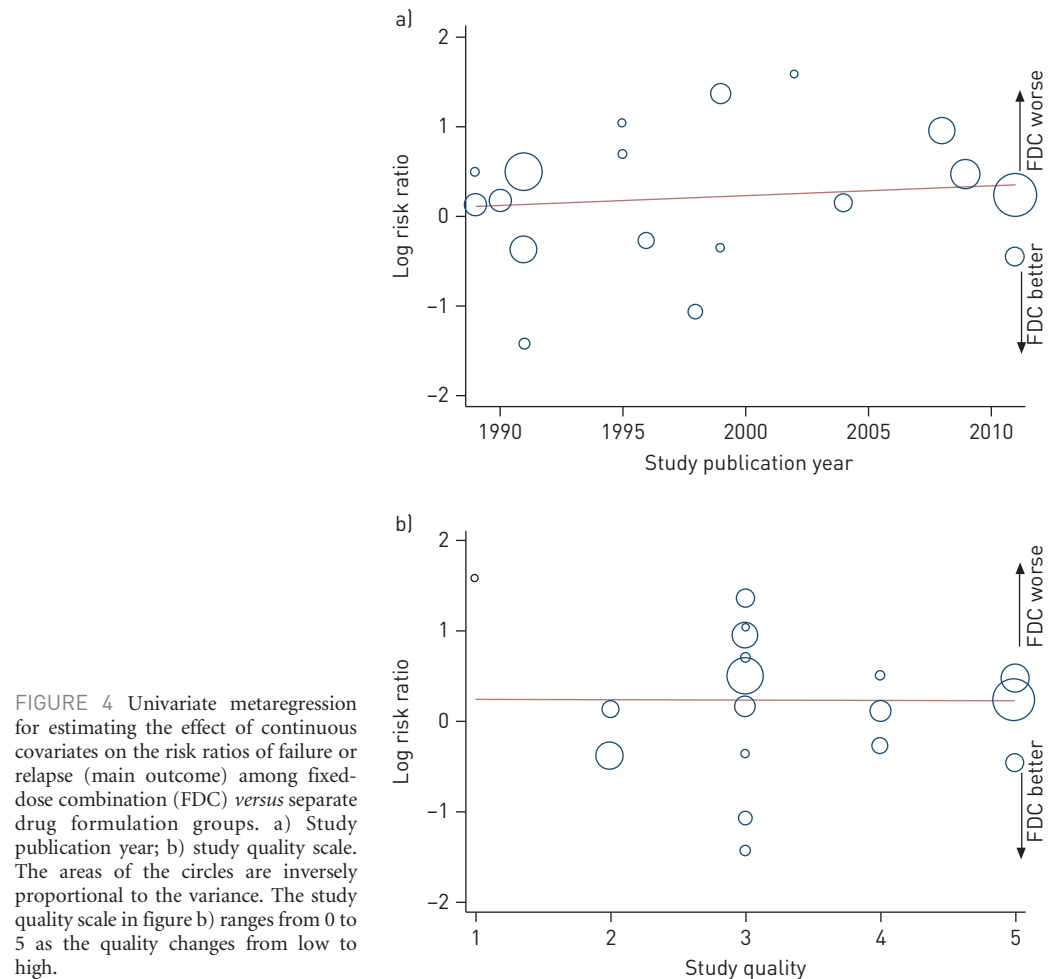


FIGURE 4 Univariate metaregression for estimating the effect of continuous covariates on the risk ratios of failure or relapse (main outcome) among fixed-dose combination (FDC) versus separate drug formulation groups. a) Study publication year; b) study quality scale. The areas of the circles are inversely proportional to the variance. The study quality scale in figure b) ranges from 0 to 5 as the quality changes from low to high.

FDC treatment was only 1%, with an upper 95% CI of 2%. Using a noninferiority design, two of the included RCTs [16, 17] demonstrated a clinically insignificant risk of unfavourable outcomes with FDCs compared to separate-drug formulations. However, this study design does not address the question of whether or not FDCs improve treatment outcomes.

Despite the potential for providing the highest level of evidence in therapeutic intervention research, RCTs have been criticised because of limited generalisability of their results. RCTs are often conducted under optimal medical care and may underestimate the potential benefit of using FDC formulations to enhance adherence in settings where malpractice or unmonitored therapies are common. In spite of this limitation, however, important differences in adherence have been found in many randomised trials [44]. To better estimate treatment effectiveness, pragmatic clinical trials may be more appropriate as these trials are conducted in a way that more closely resembles usual clinical practice [45, 46].

We designed our research protocol to include observational studies, despite their inherent susceptibility to confounding, since they better reflect real medical practice. However, only one comparative cohort study [39], which presented crude estimates that were not adjusted for potential confounding, met the inclusion criteria. Failure to adjust for potential confounding in this observational study may have reduced the validity of results, since the use of FDC formulations may correlate with adherence to other standard treatment recommendations that influence disease outcomes. Because of this limitation and because the results of this comparative cohort were significantly different from the RCT results, we did not pool both results.

One of the limitations of this meta-analysis is the small number of studies that investigated the risk of acquired drug resistance, resulting in less precise estimates. Another limitation is the inconsistent ascertainment methods of patient adherence and treatment satisfaction in different studies; because of these heterogeneous methods, we did not pool these study results. In addition, we could not assess mortality as an

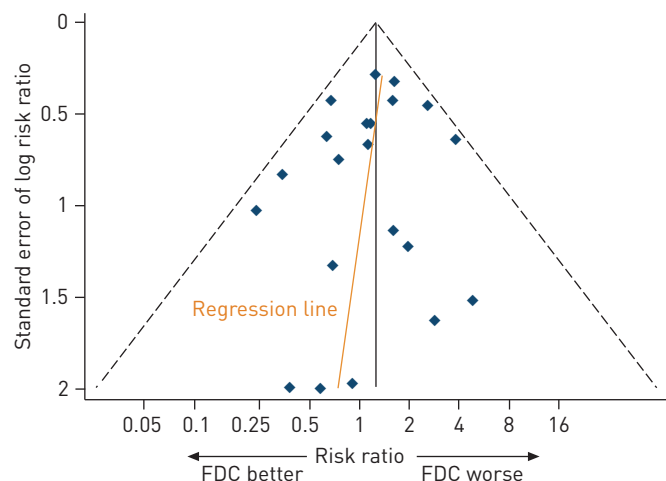


FIGURE 5 Funnel plot for the “treatment failure or disease relapse” outcome with pseudo 95% confidence limits. FDC: fixed-dose combination. Egger’s regression line represents the effect of smaller studies (higher standard error) as compared to the larger studies (lower standard error).

outcome because it was defined differently in the studies (all-cause *versus* TB-specific mortality), measured over different follow-up periods, ranging from 1 to 5 years, and in some studies was not reported or was not attributed to treatment group.

Despite these limitations, this systematic review has a number of strengths. Our systematic review was conducted without language restriction to accurately represent the existing evidence. Lack of significant heterogeneity of the estimates of treatment failure or disease relapse in the different trials permitted pooling and increased precision of our results. Another strength is the ability to stratify subjects based on their baseline drug susceptibility, which was a significant covariate factor influencing the risk of treatment failure or disease relapse. Comparing FDC to separate-drug formulation treatments, this risk tended to be higher within the stratum of subjects with baseline drug-susceptible TB and lower (in favour of FDC) within the stratum of subjects with baseline drug-resistant TB. This finding was unexpected because FDC formulations, which contain first-line anti-TB drugs, are inappropriate for patients with disease that is resistant to one or more of its component drugs. However, the result of the drug-resistant stratum included small numbers of patients with very heterogeneous forms of resistance to anti-TB drugs.

In summary, we used a strict search strategy to limit subjective selection of published studies; combined study results only when appropriate, using random effect meta-analysis which accounts for between-study variations; and followed the preferred reporting items for systematic reviews and meta-analyses (PRISMA) Statement [47] to report our data. Despite the advantage of FDC formulations in simplifying drug supply management (procurement, storage and distribution), doctor’s prescription and patient consumption of anti-TB medications, this systematic review provides evidence that FDC formulations are not superior to separate-drug formulations for preventing treatment failure or disease relapse. Furthermore, there is no evidence that FDC formulations will improve patient compliance, and inconsistent evidence that FDC regimens improve treatment satisfaction. These findings may not be generalisable to settings with unstandardised or uncontrolled medical practice.

This systematic review of current evidence does not support the use of FDC formulations for the purpose of improving treatment outcomes among patients with active TB. To provide high-quality evidence for health policies and clinical decisions, further research on clinical effectiveness of FDC anti-TB formulations should utilise pragmatic trial designs to simulate real-world clinical practice while minimising confounding.

### Acknowledgements

The authors thank Madhukar Pai (Dept of Epidemiology and Biostatistics, McGill University, Montreal, Canada) and Andrea Benedetti (Montreal Chest Institute, McGill University) for assistance in data analysis; Ioana Nicolau and Daniel Magin (McGill University), Idun Liu (Montreal Chest Institute) and Antje Bier (Royal Victoria Hospital, McGill University) for support in translation of foreign language articles; and Christian Lienhardt (Stop TB Partnership, World Health Organization, Geneva, Switzerland) and Otto Braendli (Swiss Lung Foundation, Wald, Switzerland) for providing additional information related to their studies.

### References

- 1 World Health Organization (WHO). Global tuberculosis control: WHO report 2012. WHO/HTM/TB/2012.6.
- 2 World Health Organization (WHO). Multidrug and extensively drug-resistant TB (M/XDR-TB): 2010 global report on surveillance and response. WHO/HTM/TB/2010.3.
- 3 Mitchison DA. How drug resistance emerges as a result of poor compliance during short course chemotherapy for tuberculosis. *Int J Tuberc Lung Dis* 1998; 2: 10–15.

- 4 van Altena R, van Soolingen D, van der Werf TS. Isoniazid resistant TB and non-compliance. *Thorax* 2004; 59: 1098.
- 5 Norval PY, Blomberg B, Kitler ME, *et al.* Estimate of the global market for rifampicin-containing fixed-dose combination tablets. *Int J Tuberc Lung Dis* 1999; 3: Suppl. 3, S292–S300.
- 6 Bangalore S, Kamalakkannan G, Parkar S, *et al.* Fixed-dose combinations improve medication compliance: a meta-analysis. *Am J Med* 2007; 120: 713–719.
- 7 The promise and reality of fixed dose combinations with rifampicin. A joint statement of the International Union Against Tuberculosis and Lung Disease and the Tuberculosis Programme of the World Health Organization. *Tuber Lung Dis* 1994; 75: 180–181.
- 8 Shishoo CJ, Shah SA, Rathod IS, *et al.* Impaired bioavailability of rifampicin in presence of isoniazid from fixed dose combination (FDC) formulation. *Int J Pharm* 2001; 228: 53–67.
- 9 Pillai G, Fourie PB, Padayatchi N, *et al.* Recent bioequivalence studies on fixed-dose combination anti-tuberculosis drug formulations available on the global market. *Int J Tuberc Lung Dis* 1999; 3: Suppl. 3, S309–S316.
- 10 Milán-Segovia RC, Domínguez-Ramírez AM, Jung-Cook H, *et al.* Relative bioavailability of rifampicin in a three-drug fixed-dose combination formulation. *Int J Tuberc Lung Dis* 2010; 14: 1454–1460.
- 11 Singh S, Mariappan TT, Shankar R, *et al.* A critical review of the probable reasons for the poor variable bioavailability of rifampicin from anti-tubercular fixed-dose combination (FDC) products, and the likely solutions to the problem. *Int J Pharm* 2001; 228: 5–17.
- 12 Assuring bioavailability of fixed-dose combinations of anti-tuberculosis medications. A joint statement of the International Union Against Tuberculosis and Lung Disease and the World Health Organization. *Int J Tuberc Lung Dis* 1999; 3: Suppl. 3, S282–S283.
- 13 WHO Model List of Essential Medicines. 17th list (March 2011). WHO Drug Information 2011.
- 14 Moulding TS, Le HQ, Rikleen D, *et al.* Preventing drug-resistant tuberculosis with a fixed dose combination of isoniazid and rifampin. *Int J Tuberc Lung Dis* 2004; 8: 743–748.
- 15 Suryanto AA, van den Broek J, Hatta M, *et al.* Is there an increased risk of TB relapse in patients treated with fixed-dose combination drugs in Indonesia? *Int J Tuberc Lung Dis* 2008; 12: 174–179.
- 16 Lienhardt C, Cook SV, Burgos M, *et al.* Efficacy and safety of a 4-drug fixed-dose combination regimen compared with separate drugs for treatment of pulmonary tuberculosis: the Study C randomized controlled trial. *JAMA* 2011; 305: 1415–1423.
- 17 Bartacek A, Schütt D, Panosch B, *et al.* Comparison of a four-drug fixed-dose combination regimen with a single tablet regimen in smear-positive pulmonary tuberculosis. *Int J Tuberc Lung Dis* 2009; 13: 760–766.
- 18 World Health Organization (WHO). Treatment of tuberculosis: guidelines. 4th Edn. (WHO/HTM/TB/2009.420). Geneva, World Health Organization, 2009.
- 19 International Standards for Tuberculosis Care (ISTC). 2nd Edn. The Hague, Tuberculosis Coalition for Technical Assistance, 2009.
- 20 American Thoracic Society, CDC, Infectious Diseases Society of America. Treatment of tuberculosis. *MMWR Recomm Rep* 2003; 52: 1–77.
- 21 Higgins JPT, Green S, eds. Cochrane Handbook for Systematic Reviews of Interventions. Chichester, The Cochrane Collaboration, 2011.
- 22 Research Committee of the Tuberculosis Association of India. Fifth tuberculosis association of India: short course chemotherapy trial. *Ind J Tuberc* 1989; 36: 95–101.
- 23 Cowie RL, Brink BA. Short-course chemotherapy for pulmonary tuberculosis with a rifampicin-isoniazid-pyrazinamide combination tablet. *S Afr Med J* 1990; 77: 390–391.
- 24 Controlled trial of 2, 4, and 6 months of pyrazinamide in 6-month, three-times-weekly regimens for smear-positive pulmonary tuberculosis, including an assessment of a combined preparation of isoniazid, rifampin, and pyrazinamide. Results at 30 months. Hong Kong Chest Service/British Medical Research Council. *Am Rev Respir Dis* 1991; 143: 700–706.
- 25 Acceptability, compliance, and adverse reactions when isoniazid, rifampin, and pyrazinamide are given as a combined formulation or separately during three-times-weekly antituberculosis chemotherapy. Hong Kong Chest Service/British Medical Research Council. *Am Rev Respir Dis* 1989; 140: 1618–1622.
- 26 Glatthaar E, Summers FS, Carlier ND. A comparative community-based therapy trial with a single combination product (Rifater-80). *J Compr Health* 1991; 2: 155–158.
- 27 Macnab MF, Bohmer PD, Seager JR. Evaluation of the 3-drug combination, Rifater, versus 4-drug therapy in the ambulatory treatment of tuberculosis in Cape Town. *S Afr Med J* 1994; 84: 325–328.
- 28 Chaulet P, Boulahbal F. Essai clinique d'une combinaison en proportions fixes de trios médicaments dans le traitement de la tuberculose. [Clinical trial of a combination of three drugs in fixed proportions in the treatment of tuberculosis. Groupe de Travail sur la Chimiothérapie de la Tuberculose.] *Tuber Lung Dis* 1995; 76: 407–412.
- 29 Agounitane D, Chiheb M, Khaled S, *et al.* Essai thérapeutique d'une combinaison de trois médicaments essentiels dans la chimiothérapie courte de la tuberculose. Resultats six mois après la fin du traitement. [A therapeutic trial of a combination of 3 essential drugs in a short course of chemotherapy in tuberculosis. Results 6 months after the end of treatment.] *Rev Mal Respir* 1990; 7: 209–213.
- 30 Bellabas M, Khaled S, Khaled NA, *et al.* Essai thérapeutique d'une combinaison d'isoniazide, de rifampicine et de pyrazinamide au cours des deux premiers mois du traitement de la tuberculose pulmonaire. [Therapeutic trial of a combination of isoniazid, rifampicin and pyrazinamide in the first 2 months of treatment of pulmonary tuberculosis.] *Rev Mal Respir* 1989; 6: 59–64.
- 31 Zhang LX, Kan GQ, Tu DH, *et al.* Fixed-dose combination chemotherapy versus multiple, single-drug chemotherapy for tuberculosis. *Curr Ther Res* 1996; 57: 849–856.
- 32 Zhu L, Yan B, Ma W. [Controlled clinical study on efficacy of fixed-dose compounds rifater/rifinah in antituberculous chemotherapy.] *Zhonghua Jie He He Hu Xi Za Zhi* 1998; 21: 645–647.
- 33 Teo SK. Assessment of a combined preparation of isoniazid, rifampicin and pyrazinamide (Rifater) in the initial phase of chemotherapy in three 6-month regimens for smear-positive pulmonary tuberculosis: a five-year follow-up report. *Int J Tuberc Lung Dis* 1999; 3: 126–132.
- 34 Assessment of a daily combined preparation of isoniazid, rifampin, and pyrazinamide in a controlled trial of three 6-month regimens for smear-positive pulmonary tuberculosis. Singapore Tuberculosis Service/British Medical Research Council. *Am Rev Respir Dis* 1991; 143: 707–712.

- 35 Su WJ, Perng RP. Fixed-dose combination chemotherapy (Rifater/Rifinah) for active pulmonary tuberculosis in Taiwan: a two-year follow-up. *Int J Tuberc Lung Dis* 2002; 6: 1029–1032.
- 36 Munteanu I, Husar I, Didilescu C, *et al.* Consideratii privind eficacitatea regimurilor terapeutice ce au in componentă combinatii fixe Rifampicină-Isoniazidă in tratamentul tuberculozei pulmonare. [Considerations about the efficiency of treatment regimens with fixed Rifampicin-Isoniazid combinations in pulmonary tuberculosis.] *Pneumologia* 2004; 53: 23–25.
- 37 Xu W-G, Lu W, He H-J, *et al.* [Randomized control study on domestic fixed-dose combinations in the initial treatment of smears positive tuberculosis.] *Zhonghua Jie He He Hu Xi Za Zhi* 2004; 27: 690–693.
- 38 Gravendeel JM, Asapa AS, Becx-Bleumink M, *et al.* Preliminary results of an operational field study to compare side-effects, complaints and treatment results of a single-drug short-course regimen with a four-drug fixed-dose combination (4FDC) regimen in South Sulawesi, Republic of Indonesia. *Tuberculosis* 2003; 83: 183–186.
- 39 Sokolova GB, Semenova OV, Bogadel'nikova IV, *et al.* Mairin-P v kompleksnoi terapii tuberkuleza. [Myrin P in the combined treatment of tuberculosis.] *Antibiot Khimioter* 2002; 47: 22–27.
- 40 Brändli O, Dreher D, Morger D. Ergebnisse der Tuberkulose-Kurzzeittherapie mit Isoniazid, Rifampicin und Pyrazinamid. [Results of short-term tuberculosis therapy with isoniazid, rifampicin and pyrazinamide.] *Schweiz Med Wochenschr* 1993; 123: 1300–1306.
- 41 Brändli O, Haegi V, Villiger B, *et al.* Kurzzeittherapie der Lungentuberkulose mit einer fixen Kombination von Isoniazid, Rifampicin und Pyrazinamid. Verlaufsbeobachtung über 2 Jahre. [Short-term therapy of lung tuberculosis using a fixed combination of isoniazid, rifampicin and pyrazinamide. Results after 2 years.] *Schweiz Med Wochenschr* 1989; 119: 299–305.
- 42 Punnotok J, Pumprueg U, Chakorn TA. A comparison of two short course tuberculosis chemotherapy regimens, both using Rifater during an intensive phase, with a 3 year follow-up. *J Med Assoc Thai* 1995; 78: 298–304.
- 43 Churchyard GJ, Corbett EL, Kleinschmidt I, *et al.* Drug-resistant tuberculosis in South African gold miners: incidence and associated factors. *Int J Tuberc Lung Dis* 2000; 4: 433–440.
- 44 Haynes RB, Ackloo E, Sahota N, *et al.* Interventions for enhancing medication adherence. *Cochrane Database Syst Rev* 2008; 2: CD000011.
- 45 Chalkidou K, Tunis S, Whicher D, *et al.* The role for pragmatic randomized controlled trials (pRCTs) in comparative effectiveness research. *Clin Trials* 2012; 9: 436–446.
- 46 Tunis SR, Stryer DB, Clancy CM. Practical clinical trials: increasing the value of clinical research for decision making in clinical and health policy. *JAMA* 2003; 290: 1624–1632.
- 47 Moher D, Liberati A, Tetzlaff J, *et al.* Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *PLoS Med* 2009; 6: e1000097.