

Sample sizes for clinical trials using sputum eosinophils as a primary outcome

Angira Dasgupta¹, Shiyuan Zhang^{2,3}, Lehana Thabane^{2,3} and Parameswaran Nair¹

Affiliations:

¹Dept of Medicine, St Joseph's Healthcare and McMaster University, Hamilton, ON, ²Dept of Clinical Epidemiology and Biostatistics, McMaster University, Hamilton, ON, and ³Biostatistics Unit, St Joseph's Healthcare, Hamilton, ON, Canada.

Correspondence:

P. Nair, Firestone Institute for Respiratory Health, St Joseph's Healthcare, 50 Charlton Avenue East, Hamilton, Ontario, L8N4A6, Canada. E-mail: parames@mcmaster.ca

ABSTRACT Clinical trials do not report sputum eosinophil data in a consistent method and this makes it difficult to compare across studies and to evaluate the sample sizes estimated in these studies. The objectives of the paper are: 1) to systematically review reporting of effect size and sample calculations in randomised controlled trials using sputum eosinophil count as a primary outcome and 2) to illustrate sample size estimation under different methods of data representation using data from an effective anti-eosinophil treatment strategy (mepolizumab).

Randomised controlled trials in adults (excluding allergen provocation models) of treatment of asthma and chronic obstructive pulmonary disease for the past 10 years were searched in Ovid MEDLINE and 20 studies were identified that met all the inclusion criteria. Only nine studies discussed sample size calculation.

Change from baseline was used as an outcome in 11 studies and was expressed as change in absolute percentage count, percentage change from baseline or as fold changes.

Assuming a minimal clinically important reduction of 15% in absolute terms, 18 subjects in each arm will be required to achieve 80% power using an ANCOVA analysis, which we recommend, to detect significance with an alpha error of 0.05.



@ERSpublications

Systematic review and illustration of sample size calculations in RCTs using sputum eosinophil count as a primary outcome $\frac{1}{M} = \frac{1}{M} \frac{1}{M} = \frac{1}{M} \frac{1}{M} \frac{1}{M} = \frac{1}{M} \frac{1}{M} \frac{1}{M} \frac{1}{M} = \frac{1}{M} \frac{$

For editorial comments see page 891.

Received: May 11 2012 | Accepted after revision: Dec 01 2012 | First published online: Jan 24 2013

Support statement: P. Nair is supported by a Canada Research Chair in Airway Inflammometry.

Conflict of interest: Disclosures can be found alongside the online version of this article at www.erj.ersjournals.com

Copyright ©ERS 2013

Introduction

Eosinophil counts in sputum are a valid and reliable outcome measure in clinical trials of patients with eosinophilic lung diseases like asthma [1, 2]. Clinical trials of corticosteroids and specific anti-eosinophil agents, like mepolizumab, have consistently shown a reduction in sputum eosinophil counts [3–6]. The recent guidelines for clinical end-points in asthma trials set out by the American Thoracic Society (ATS) and the European Respiratory Society (ERS) have also incorporated the use of induced sputum eosinophil counts as an outcome measure [7]. However, this document does not provide guidelines for expressing results or calculating sample sizes for clinical trials which use sputum eosinophils as an outcome.

Clinical trials with two groups and with pre- and post-interventional data in each group usually compare either the post-interventional scores or change from baseline scores, using either a t-test or an analysis of covariance adjusting for baseline differences. The latter method of analysis has been held to be superior to all other methods by statisticians, as it has the highest statistical power [8]. For sputum eosinophil counts, which are conventionally expressed as a percentage of the total cell count, change from baseline is preferred over post-interventional scores as it is clinically more meaningful and gives an estimate of the amount of improvement or deterioration caused by an intervention. It has been observed that clinical trials reporting change from baseline for sputum eosinophil counts have expressed the results as fold changes (ratio of preand post-interventional scores), percentage from baseline or as change in absolute percentage counts. Although none of the three are superior to one another from a biological point of view, change from baseline in terms of absolute percentage count is preferable due to its better clinical interpretability and lack of need for statistical data manipulation. Expressing results as fold changes or percentage change from baseline may be challenging, especially when there are zero counts as often happens in clinical practice. These methods also convert the values into ratios, which have a non-normal or skewed distribution necessitating data transformation (usually log) prior to carrying out statistical tests that assume a normal distribution. The zero values are usually either replaced by positive number, like 0.1 or 0.001 [9], or a constant number, e.g. 0.2 [7], added prior to log transformation.

As a consequence of the differences in expressing the results, sample size estimation methods vary significantly in these clinical trials, as sample sizes depend on the selected effect sizes and the standard deviation of the primary outcome variable [10]. Often sample size reporting is also incomplete. We are not aware of any study, so far, that has examined the completeness or the quality of reporting of sample size calculations, effect sizes and analysis methods, or demonstrated the method of doing so for clinical trials using change in sputum eosinophil count from baseline as the primary outcome.

We therefore decided to 1) systematically review current practice of reporting effect sizes and sample size calculations in randomised controlled trials using sputum eosinophil count as a primary outcome; 2) describe and illustrate how to apply different strategies for defining outcomes using sputum eosinophil count in interventional studies that have both baseline and post-intervention data; and 3) demonstrate sample size estimation under the different methods of data representation illustrated in 2) and data analysis with a dataset from a prior clinical trial [3].

Methods

Search strategy

The OVID search engine was used to search for all randomised controlled trials, in the English language, in the past 10 years, with "asthma" AND "sputum eosinophils" OR "airway inflammation" as key search words using the Ovid MEDLINE(R), EMBASE and Ovid HealthSTAR databases. The inclusion criteria of the articles were: 1) studies on asthma and chronic obstructive pulmonary diseases (COPD); 2) randomised controlled trials involving drug trials on human subjects; and 3) studies using sputum eosinophil measures as a primary outcome. Both parallel group and cross over studies were included. Studies were included if the primary aim was to look at airway inflammation even if sputum eosinophil count was not the only primary outcome. Studies involving children, allergen or antigen provocation models, environmental exposure studies and those using nonparametric tests as the primary methods of analysis were excluded. Data were extracted independently by two reviewers (A. Dasgupta and S.Z.) after going through the methods section of all the studies and tabulated carefully using a data extraction form. Any disagreements between the two reviewers were discussed and resolved by consensus. Figure 1 shows the study selection process.

Different strategies for defining outcomes

Sputum eosinophil counts are customarily expressed as percentages of total cell count and not as absolute counts [11]. Change from baseline for sputum eosinophils for asthma and COPD trials can be represented in three different ways as follows: 1) change from baseline in absolute percentage counts, *i.e.* pre-intervention % - post-intervention % (or post-intervention % - pre-intervention %); 2) fold changes, *i.e.*

1004 DOI: 10.1183/09031936.00075712

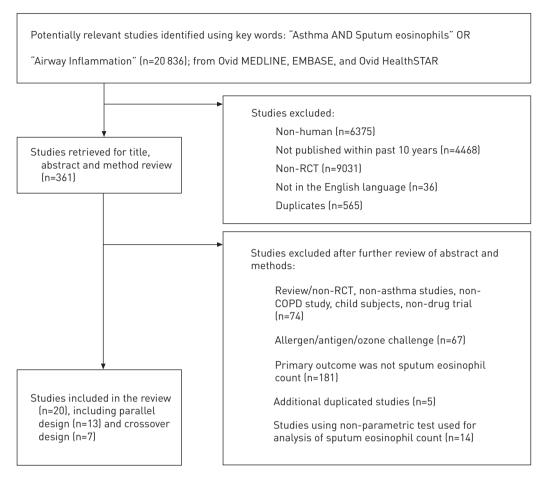


FIGURE 1 Study selection. The data in this figure refer to the search completed on March 15, 2012. RCT: randomised controlled trials; COPD: chronic obstructive pulmonary disease.

pre-intervention % / post-intervention % (or post-intervention % / pre-intervention %); 3) percentage change from baseline, *i.e.* (pre-intervention % – post-intervention %) \times 100/pre-intervention %.

Data from the mepolizumab study [3] was used to demonstrate the above strategies. Zero values were replaced by 0.1 prior to all calculations.

Demonstration of sample size calculation

Sample size calculations were demonstrated using the above dataset for independent parallel group superiority studies for t-tests to achieve 80% power and a Type I error of 0.05 assuming equal sample sizes in intervention and placebo group. The standard deviations and the effect sizes for change in absolute percentage counts were calculated. Fold changes were log transformed (natural logs) and zero counts replaced by 0.1 prior to transformation. We also used the data to calculate the sample size with fold changes as the outcome. For this calculation, both standard deviation and effect estimate were on log scale.

The sample size formula [10] used was:

$$n = 2 \times s^2 \times \frac{\left(Z_{\alpha/2} + Z_{1-\beta}\right)^2}{\delta_0^2}$$

Where n is the size per group; $Z\alpha/2$ is the standard normal z-score corresponding to the two-sided α level of significance; Z_1 - β is the standard normal z-score corresponding to the probability β of Type II error; δ_0 is the minimum clinically important difference; s is the prior estimate of the pooled standard deviation of the estimate of the difference between groups.

TABLE 1 Details of method used for effect representation and samples sizes for the 20 included studies

First author [ref.]	Journal	Year of publication	Method of data representation	Type of analysis for between groups	Total sample size	Study design
GREEN [14]	Lancet	2002	Log-transformed, Percentage difference between groups	Post score comparison by t-test	74	Parallel
KANNIESS [15]	European Respiratory Journal	2002	Fold changes	Change from baseline by ANOVA	40	Crossover
MINOGUCHI [16]	Chest	2002	Absolute percentage counts	Change from baseline by ANOVA	29	Crossover
URATA [17]	Respiratory Medicine	2002	Absolute percentage counts	Post score comparison by t-test	33	Crossover
Yoshida [18]	Journal of Asthma	2002	Absolute percentage counts	Post score comparison by t-test	28	Parallel
GROOTENDORST [19]	Pulmonary Pharmacology and Therapeutics	2003	Log-transformed	Post scores comparison by ANOVA	18	Crossover
Djukanovíć [20]	American Journal of Respiratory and Critical Care Medicine	2004	Absolute percentage counts	Post score comparison by ANCOVA	45	Parallel
KOBAYASHI [21]	Respirology	2004	Percentage change from baseline	Change from baseline by repeated measures ANOVA	11	Parallel
TEN BRINKE [22]	American Journal of Respiratory and Critical Care Medicine	2004	Log-transformed percentage count	Post score comparison by ANCOVA	22	Parallel
JAYARAM [23]	European Respiratory Journal	2005	Log transformed, percentage difference in treatment between groups	Post scores comparison by ANCOVA	14	Crossover
JAYARAM [13]	Thorax	2005	Log-transformed, absolute percentage counts and per- centage change from base- line	Change from baseline by ANCOVA	50	Crossover
GREEN [24]	European Respiratory Journal	2006	Log-transformed, fold change from baseline	Change from baseline by two-way ANOVA	49	Crossover
HAUBER [25]	Canadian Respiratory Journal	2006	Absolute percentage counts	Change from baseline by ANOVA	17	Crossover
HAVERKAMP [26]	Journal of Allergy and Clinical Immunology	2007	Absolute percentage counts	Change from baseline by unpaired t-test with Bonferroni correction	19	Parallel
PAVORD [27]	Respiratory Research	2007	Absolute percentage counts	Change from baseline by ANCOVA	66	Parallel
CHEN [28]	Pulmonary Pharmacology and Therapeutics	2008	Absolute percentage counts	Post score comparison by unpaired t-test	37	Parallel
Duong [29]	Chest	2008	Log-transformed	Change from baseline by ANOVA	26	Parallel
Ноѕнімо [30]	Allergology International	2009	Absolute percentage counts	Change from baseline by two sample t-test and Mann-Whitney U-test	27	Crossover
PAVORD [31]	Journal of Allergy and Clinical Immunology	2009	Log transformed, fold change from baseline	Change from baseline by ANCOVA	127	Parallel
DENTE [9]	Annals of Allergy, Asthma, and Immunology	2010	Log-transformed	Post score comparison by ANOVA	59	Parallel

Sample size calculation using ANCOVA was done for a hypothetical superiority trial where the clinically minimal important difference was 15% in absolute percentage change or twofold-change difference. Sample size for ANCOVA was calculated from that of the t-test while adjusting by a factor of $(1-\rho^2)$, where ρ is the correlation between the baseline and the final outcome measure [12].

Results

Reporting of sample size calculation

Of the 361 search results, 20 studies fulfilled the inclusion criteria [9, 13–32]. 11 of the 20 (55%) did not report how the sample size was calculated. Of the nine other articles that reported sample size calculation, only two articles had detailed all the required parameters related to sample size calculation and only one article contained all eight items within the checklist (table 1). The articles expressed outcome variably as change in absolute percentage counts, percentage changes from baseline or as fold changes. One article reported outcome as both change in absolute percentage count and as percentage change from baseline. Nine articles reported post-intervention score comparison, while the rest used change from baseline data for

1006 DOI: 10.1183/09031936.00075712

TABLE 2 Results of the systematic review

Item number	Descriptor	Yes	No	Unclear
1	Desired significance level	8 (40)	12 (60)	0
2	Desired power level	7 (35)	13 (65)	0
3	Type of test statistics of method of analysis for the primary outcome	7 (35)	11 (55)	2 (10)
4	Description of the primary outcome (absolute count, log-transformed or fold change)	20 (100)	0	0
5	Minimal clinically important difference	7 (35)	13 (65)	0
6	Variance for sample size estimation	2 (10)	18 (90)	0
7	Calculated sample size	7 (35)	13 (65)	0
8	Software/reference/formula for sample size calculation provided	1 (5)	17 (85)	2 (10)

Data are presented as n (%). Total number of papers extracted for analysis n=20.

analysis (table 2). Six articles used ANCOVA for statistical analysis but none of the studies adjusted sample sizes for this.

Demonstration of strategies for defining outcomes

The mepolizumab trial was a parallel group randomised control trial, with 10 patients in the placebo group and eight patients in the intervention (mepolizumab) group with pre- and post-intervention sputum eosinophil counts. The raw data for each of the groups and the different ways of expressing change from baseline are demonstrated below (table 3). On the other hand, changes in absolute percentage counts, being the difference between pre- and post-intervention values, may be assumed to be normally distributed despite the distribution of the pre- and post-interventional values being non-normal. On the other hand, fold changes being ratios have a skewed statistical distribution. Hence, fold changes were log transformed prior to calculating the standard deviation and effect estimate. Similar results for percentage change from baseline were expressed as median (range).

In the placebo group sputum eosinophils increased by mean \pm SD 17.9 \pm 27.8% while in the intervention group it decreased by 21.05 \pm 14.6%. The effect size was therefore 38.9% (absolute percentage counts) or 54-fold (or 3.8 in natural log-scale). The pooled standard deviation was 23% and 1.47 in absolute percentage counts and log (fold changes), respectively. The median change in terms of percentage change from baseline was -251.2% (-1669.4–60.6%) for the placebo group and 93.8% (71.9–99.6%) for the intervention group.

Demonstration of sample size calculation

The values of standard deviation and effect sizes expressed as change in absolute percentage counts and natural log transformed fold changes were put into equation 1 and calculations made for sample sizes (table 4). From the mepolizumab clinical trial data, the correlation between pre-intervention and change from baseline in absolute percentage counts and in the log scale is -0.73 and 0.72, respectively. Therefore, the number of participants required will be reduced by a factor of 0.47. Thus, one (2×0.47) patients in each arm or three (6×0.47) patients in each arm will be sufficient when outcomes are expressed as change in absolute percentage counts and fold changes, respectively, provided other parameters remain constant.

Discussion

Reporting sample sizes and effect sizes

Sample size calculations are under-reported in clinical drug trials using sputum eosinophil counts as a primary outcome. This is clear from the results of our systematic review. This reduces the accuracy of trials and leads to inappropriate conclusions. There are also substantial variations in the way change from baseline sputum eosinophil count is expressed, which makes comparison between clinical trials and drug effects difficult. There is, thus, an unmet need to formulate guidelines for reporting drug effects when using sputum eosinophil counts in clinical trials. We did not examine the differences in sample size estimations and reporting between asthma and COPD trials, as the variance of change of sputum eosinophil counts and the minimal clinically important difference are conventionally taken to be the same for both diseases. Therefore, sample size estimation is unlikely to be affected by the disease definition. We also did not include studies in subjects with mild intermittent asthma where effects of various interventions in attenuating allergen-induced eosinophilic inflammation were examined. This was for two reasons. First, we wanted to limit our estimation of sample sizes to clinical trials that evaluated therapies that are relevant to patients

Group	Patient	Pre-intervention#	Post-intervention#	Change from baseline in % counts ¹	Percentage change from baseline ⁺	Fold changes [§]	In of fold changes
Placebo group	_	3.7	18.7	15.0	-405.4	0.19	-1.62
	2	3.6	63.7	60.1	-1669.4	0.02	-2.87
	က	4.7	62.3	57.6	-1225.5	0.07	-2.58
	7	4.5	25.3	20.8	-462.2	0.17	-1.72
	2	4.0	5.0	1.0	-25	0.8	-0.22
	9	35.3	15.5	-19.8	56.1	2.27	0.82
	7	3.3	1.3	-2.8	9.09	2.53	0.93
	8	0.9	5.0	-1.0	16.6	1.2	0.18
	6	4.6	5.3	0.7	-15.2	98.0	-0.14
	10	3.2	50.0	8.97	-1462.5	0.064	-2.74
Mean∃sp				17.9 ± 27.8	-215.2^{f}		-1.0 ± 1.48
Mepolizumab	_	21.0	1.3	-19.7	93.8	16.15	2.78
(intervention) group	2	25.5	0.1 (0 replaced)	-25.4	9.66	255	5.54
	က	54.3	11.3	-43.0	79.1	4.80	1.57
	4	16.6	1.0	-15.6	93.9	16.6	2.81
	2	42.3	9.0	-41.7	98.5	70.5	4.26
	9	10.7	3.0	-7.7	71.9	3.57	1.27
	7	0.9	0.3	-5.7	9.5	20	2.99
	8	12.3	2.7	9.6-	78.4	4.56	1.52
Mean∃sp				-21.1 ± 14.6	93.8 [‡]		2.8 ± 1.47

 $^{\#}$: eosinophil count as a percentage of the total cell count; 1 : calculated as pre-intervention % - post-intervention %; $^{+}$: calculated as [pre-intervention %, 1 : median.

TABLE 4 Sample sizes for various representations of data using t-test

Representation of data	Means placebo [#] versus mepolizumab [¶]	Measures of dispersion	Difference between the groups, detected or relative effect	Sample sizes for each group
Change in absolute percentage count % ⁺	17.92 versus -21.05	Pooled SD=22.9	38.9	6
Mean fold change [§]	0.9 versus 48.9	Pooled sp=1.47 (In scale)	54 fold or 3.8 (In scale)	2

^{#:} n=10; ¶: n=8; +: calculated as baseline % - end of treatment %; ⁵: baseline/end of treatment.

with mild persistent or moderate-to-severe asthma in phase IIb or III studies, rather than experimental therapies in subjects with mild intermittent asthma. Secondly, and more importantly, mechanisms of reversal of established eosinophilia may be very different to those that prevent recruitment of eosinophilis (such as following an allergen inhalation) and, therefore, effect sizes for these two separate interventions may be different. This needs to be investigated in a separate study.

Expressing change from baseline

Of the three methods of expressing change from baseline, percentage change from baseline has been considered to be an inefficient method, by statisticians, as it does not correct for any imbalance between groups at baseline and has a non-normal distribution [8]. Further, when used in sputum eosinophil trials, as demonstrated with this dataset, percentage change from baseline values may be both negative and positive and also numerically very large. This makes data transformation difficult, and percentage change from baseline is, therefore, not a very useful method of data representation for sputum eosinophil studies.

By contrast, change in absolute percentage count has greater clinical interpretability compared to fold changes, as exemplified below. A change from 25% to 5% is a 20% change in absolute percentage count, which is equivalent to a fivefold change or 80% change from baseline. Again a reduction from 5 to 1% is a 4% change in absolute percentage count, but has the same fold and percentage changes as the former case. In the latter case the sputum eosinophils have been normalised and a patient with such a report will not require additional treatment. But in the former situation a patient needs to be treated further to reduce the counts into the normal range. In another example it may be shown that a change from 25% to 0.1% is a 250-fold change, which is the same fold change as a change from 5% to 0.2%. It is obvious that the former change has more clinical significance than the latter. Therefore, it may be reasonable to suggest that effect sizes or outcomes in clinical trials be expressed in terms of absolute percentage counts for better interpretation of results in the clinical context. The use of fold changes may be justified for use only in those situations where data transformation is needed for statistical purposes. In clinical practice and in

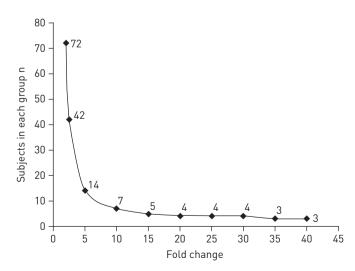


FIGURE 2 Sample sizes for demonstrating a range of fold changes in eosinophil percentage counts between the active and placebo group.

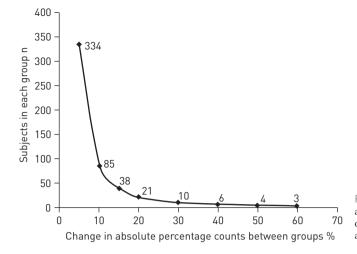


FIGURE 3 Sample sizes for demonstrating a range of absolute difference in eosin-ophil percentage counts between the active and placebo group.

dose-response studies [32], it is generally observed that sputum eosinophil percentages change proportional to the baseline values, although this has not been evaluated systematically.

Sample size calculation

Sample size calculations using the mepolizumab data show that only six subjects or fewer in each arm in a parallel group randomised study on severe asthma subjects may have sufficient power to detect significant changes when using a t-test, assuming that all participants have the same baseline risk. The reason for such small sample size estimation is the large effect size of this trial. Such large effect sizes are, however, not unusual in severe asthma trials using anti-eosinophil specific therapies [4]. Corticosteroids have shown a reduction in the range of two- to seven-fold [6, 33] or 15% to 20% in absolute sputum eosinophil percentages in clinical trials of severe eosinophilic asthma [9, 13]. Smaller effect sizes are, however, seen with milder asthma trials [14].

The standard minimal clinically important difference (MCID) for sputum eosinophil studies has been traditionally accepted to be a 2-fold or 50% change [7]. From figure 2, which shows the sample size estimates for various effect sizes, it may be appreciated that clinical trials such as the mepolizumab study will have 80% power for detecting such a change with 72 participants in each arm. A consensus agreement on the MCID in terms of absolute percentage counts has not yet been reached. However, as mentioned earlier corticosteroids have shown a reduction in eosinophil counts of between 15% and 20% in various severe asthma studies. Therefore, it may be reasonable to assume an MCID of 15% in terms of absolute percentage counts when performing severe asthma trials with baseline eosinophil counts of more than 15%. Thus, 37 and 18 subjects in each arm will be required to achieve an 80% power using a t-test and ANCOVA, respectively (fig. 3).

It may also be seen that the estimate using log transformation (in fold change calculations) is much smaller than that using absolute percentage counts. This is because the variance gets stabilised as a result of the transformation, thereby reducing the sample size. Thus, sample size estimates are also somewhat affected by the method of data representation. From our calculations there was a reduction in the sample size estimates by a factor of two when similar sample size parameters were used. However, due to the increased level of complexity when using ANCOVA, the gained statistical power is often out-balanced by the difficulty in employing the method.

Conclusion

Clinical trials using sputum eosinophil counts as a primary outcome do not often report how sample size calculations are made or the parameters assumed for them. Sample sizes in these trials are affected not only by the method of analysis, study design, MCID and variance, but also by the strategies adopted for data representation for defining outcomes. The best method to adopt for such trials is to express results as the change in absolute percentage counts, as this is clinically most relevant and does not involve the statistical intricacies of data transformation. Adjusting calculated sample sizes based on use of ANCOVA may reduce the sample size estimates by half. A sample size of 18 patients in each arm in an ANCOVA superiority trial may be sufficient to achieve an 80% power, with an alpha error of 0.05, to detect a 15% difference in absolute percentage counts in severe asthma studies with specific anti-eosinophil agents, such as the mepolizumab.

1010 DOI: 10.1183/09031936.00075712

References

- Pizzichini E, Pizzichini MM, Efthimiadis A, et al. Indices of airway inflammation in induced sputum: reproducibility and validity of cell and fluid-phase measurements. Am J Respir Crit Care Med 1996; 154: 308–317.
- Nair PK, Hargreave FE. Airway diseases, inflammometry and individualized therapy. *In*: Polosa R, Holgate ST, eds. Therapeutic Strategies. Asthma: Current Treatments. Oxford, Clinical Publishing, 2007; pp. 155–164.
- Nair P, Pizzichini MM, Kjarsgaard M, et al. Mepolizumab for prednisone-dependent asthma with sputum eosinophilia. N Engl J Med 2009; 360: 985–993.
- 4 Castro M, Mathur S, Hargreave F, et al. Reslizumab for poorly controlled, eosinophilic asthma: a randomized, placebo-controlled study. Am J Respir Crit Care Med 2011; 184: 1125–1132.
- Pavord ID, Brightling ČE, Woltmann G, et al. Non-eosinophilic corticosteroid unresponsive asthma. Lancet 1999; 353: 2213–2214.
- Pizzichini MM, Pizzichini E, Clelland L, et al. Sputum in severe exacerbations of asthma: kinetics of inflammatory indices after prednisone treatment. Am J Respir Crit Care Med 1997; 155: 1501–1508.
- 7 Reddel HK, Taylor DR, Bateman ED, et al. An official American Thoracic Society/European Respiratory Society statement: asthma control and exacerbations: standardizing endpoints for clinical asthma trials and clinical practice. Am J Respir Crit Care Med 2009; 180: 59–99.
- 8 Vickers AJ. The use of percentage change from baseline as an outcome in a controlled trial is statistically inefficient: a simulation study. *BMC Med Res Methodol* 2001; 1: 6.
- Dente FL, Bacci E, Bartoli ML, et al. Effects of oral prednisone on sputum eosinophils and cytokines in patients with severe refractory asthma. Ann Allergy Asthma Immunol 2010; 104: 464–470.
- 10 Zhong BL. How to calculate sample size in randomized controlled trial? J Thorac Dis 2009; 1: 51–54.
- Belda J, Leigh R, Parameswaran K, et al. Induced sputum cell counts in healthy adults. Am J Respir Crit Care Med 2000; 161: 475–478.
- 12 Borm GF, Fransen J, Lemmens WA. A simple sample size formula for analysis of covariance in randomized clinical trials. J Clin Epidemiol 2007; 60: 1234–1238.
- 13 Jayaram L, Pizzichini E, Lemière C, et al. Steroid naive eosinophilic asthma: anti-inflammatory effects of fluticasone and montelukast. Thorax 2005; 60: 100–105.
- 14 Green RH, Brightling CE, McKenna S, et al. Asthma exacerbations and sputum eosinophil counts: a randomised controlled trial. Lancet 2002; 360: 1715–1721.
- 15 Kanniess F, Richter K, Böhme S, et al. Montelukast versus fluticasone: effects on lung function, airway responsiveness and inflammation in moderate asthma. Eur Respir J 2002; 20: 853–858.
- Minoguchi K, Kohno Y, Minoguchi H, et al. Reduction of eosinophilic inflammation in the airways of patients with asthma using montelukast. Chest 2002; 121: 732–738.
- 17 Urata Y, Yoshida S, Irie Y, *et al.* Treatment of asthma patients with herbal medicine TJ-96: a randomized controlled trial. *Respir Med* 2002; 96: 469–474.
- 18 Yoshida M, Aizawa H, Inoue H, *et al.* Effect of suplatast tosilate on airway hyperresponsiveness and inflammation in asthma patients. *J Asthma* 2002; 39: 545–552.
- Grootendorst DC, Gauw SA, Benschop N, et al. Efficacy of the novel phosphodiesterase-4 inhibitor BAY 19-8004 on lung function and airway inflammation in asthma and chronic obstructive pulmonary disease (COPD). Pulm Pharmacol Ther 2003; 16: 341–347.
- 20 Djukanović R, Wilson SJ, Kraft M, et al. Effects of treatment with anti-immunoglobulin E antibody omalizumab on airway inflammation in allergic asthma. Am J Respir Crit Care Med 2004; 170: 583–593.
- 21 Kobayashi M, Nasuhara Y, Betsuyaku T, et al. Effect of low-dose theophylline on airway inflammation in COPD. Respirology 2004; 9: 249–254.
- ten Brinke A, Zwinderman AH, Sterk PJ, et al. "Refractory" eosinophilic airway inflammation in severe asthma: effect of parenteral corticosteroids. Am J Respir Crit Care Med 2004; 170: 601–605.
- 23 Jayaram L, Duong M, Pizzichini MM, et al. Failure of montelukast to reduce sputum eosinophilia in high-dose corticosteroid-dependent asthma. Eur Respir J 2005; 25: 41–46.
- 24 Green RH, Brightling CE, McKenna S, et al. Comparison of asthma treatment given in addition to inhaled corticosteroids on airway inflammation and responsiveness. Eur Respir J 2006; 27: 1144–1151.
- 25 Hauber H, Taha R, Bergeron C, et al. Effects of hydrofluoroalkane and dry powder-formulated corticosteroids on sputum inflammatory markers in asthmatic patients. Can Respir J 2006; 13: 73–78.
- Haverkamp HC, Dempsey JA, Pegelow DF, et al. Treatment of airway inflammation improves exercise pulmonary gas exchange and performance in asthmatic subjects. J Allergy Clin Immunol 2007; 120: 39–47.
- Pavord I, Woodcock A, Parker D, et al. Salmeterol plus fluticasone propionate versus fluticasone propionate plus montelukast: a randomised controlled trial investigating the effects on airway inflammation in asthma. Respir Res 2007; 8: 67.
- 28 Chen YH, Yao WZ, Ding YL, et al. Effect of theophylline on endogenous hydrogen sulfide production in patients with COPD. Pulm Pharmacol Ther 2008; 21: 40–46.
- Duong M, Subbarao P, Adelroth E, *et al.* Sputum eosinophils and the response of exercise-induced bronchoconstriction to corticosteroid in asthma. *Chest* 2008; 133: 404–411.
- 30 Hoshino M, Handa H, Miyazawa T. Effects of salmeterol and fluticasone propionate combination versus fluticasone propionate on airway function and eosinophilic inflammation in mild asthma. Allergol Int 2009; 58: 357–363.
- Pavord ID, Jeffery PK, Qiu Y, et al. Airway inflammation in patients with asthma with high-fixed or low-fixed plus as-needed budesonide/formoterol. J Allergy Clin Immunol 2009; 123: 1083–1089.
- 32 Kelly MM, Leigh R, Jayaram L, et al. Eosinophilic bronchitis in asthma: a model for establishing dose-response and relative potency of inhaled corticosteroids. J Allergy Clin Immunol 2006; 117: 989–994.
- Turner MO, Johnston PR, Pizzichini E, et al. Anti-inflammatory effects of salmeterol compared with beclomethasone in eosinophilic mild exacerbations of asthma: a randomized, placebo controlled trial. Can Respir J 1998; 5: 261–268.