Long- and medium-term ozone effects on lung growth including a broad spectrum of exposure

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Long- and medium-term ozone effects on lung growth including a broad spectrum of exposure. G. Ihorst, T. Frischer, F. Horak, M. Schumacher, M. Kopp, J. Forster, J. Mattes, J. Kuehr. ©ERS Journals Ltd 2004.

ABSTRACT: The effects of semi-annual and 3.5 yr mean ozone (O_3) concentrations on children's forced vital capacity (FVC) and forced expiratory volume in one second (FEV1) were assessed over a study period of 3.5 yrs in 2,153 schoolchildren from 15 study sites in South Western Germany and Lower Austria.

Spirometric parameters were assessed twice a year, and differences between consecutive measurements divided by days were considered as a measure of lung growth. Exposure was analysed in four classes, separately for winter and summer (semi-annual mean O₃ concentrations 22–30, 30–38, 38–46, 46–54 parts per billion (ppb) in summer and 4–12, 12–20, 20–28, 28–36 ppb in winter).

Regression methods for repeated measurements were used, and these revealed a significantly lower FVC (FEV1) increase estimated at -19.2 (-18.5) mL·100 days $^{-1}$ for semi-annual mean O_3 exposure in summer between 46 and 54 ppb compared to exposure between 22 and 30 ppb. However, in winter, the estimated difference in FVC (FEV1) was 16.4 (10.9) mL·100 days $^{-1}$ between the semi-annual O_3 class 28–36 ppb and the 4–12 ppb class. By means of linear regression the study found that there was no association between growth rates and mean summer O_3 for FVC and FEV1 over a 3.5 yr period.

The authors conclude that medium-term effects on schoolchildren's lung growth are possibly present, but are in the long-term not detectable for forced vital capacity and forced expiratory volume in one second over a 3.5 yr period due to partial reversibility. Eur Respir J 2004; 23: 292–299.

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Various studies have demonstrated the effects of short-term ozone (O_3) exposure on children's respiratory health or lung function parameters such as forced vital capacity (FVC), forced expiratory volume in one second (FEV1) or peak expiratory flow (PEF). The expression short-term exposure refers to time spans of 1 h to several days before lung function is tested. Medium-term exposure means time periods of several months and long-term exposure refers to years (without representation of seasonal effects). Children are considered to be at higher risk than adults to possible damage caused by air pollution, since they spend more time outdoors, are more active and their growing lungs might be more susceptible.

Some of these studies were conducted under conditions that were natural but with high exposure values, *i.e.* lung function measurements were made before and after physical exercise [1–3]. Some studies took place in summer camps where children spent much time outside and were thus exposed to air pollution most of the time. A combined analysis of six summer camp studies showed a significant negative effect of O₃ on FEV1 but not on PEF [4]. Other investigators studying children during normal daily activities found negative associations between O₃ and lung function measurements [5–13].

To the best of the authors' knowledge only a few longitudinal investigations have been made concerning the longer-lasting effects of O₃. A preliminary analysis of the first

cross section of a study with schoolchildren from 12 South California communities suggested negative associations of chronic O₃ exposure and expiratory flow rates [14]. During the follow-up part of the study no further O₃ effect could be determined [15]. The same study group investigating a different cohort over a 4-yr period observed an O₃ effect on peak expiratory flow rate (PEFR) [16]. KINNEY *et al.* [17] found a seasonal decline in respiratory function in healthy young adults working outside over one summer period in the presence of O₃ and particulate matter. Associations between ambient O₃ exposure and lung growth in schoolchildren were reported in studies from Austria [18].

In order to investigate chronic effects of inhaled air pollutant on lung growth repeated measurements on each child are required [19, 20]. Therefore two co-conducted studies were carried out in 15 communities in Germany (Black Forest) and Lower Austria with different levels of ambient O₃ concentrations, performing repeated spirometry measurements on elementary schoolchildren twice a year (in spring and autumn). Earlier results of the first three study years of the Austrian study alone [18] and a combined analysis of Austrian data from the first 2 yrs together with an earlier German study [21] had already shown signs of delayed lung growth in summer and a catch-up in winter for children from the regions with higher seasonal O₃ exposure. The current authors are now able to present data from a 4-yr time period for geographical areas representing a large spectrum of outdoor O₃ concentrations.

In total 6,814 spirometries were obtained from 1,114 Austrian children and 6,105 spirometries from 1,039 German children. The difference between 2 measurements divided by time is considered to be an adequate description of lung growth. The current article focuses on two questions. Does high O₃ exposure result in decreased lung growth over a period up to 3.5 yrs (long-term)? Does children's lung growth vary with seasonal exposure to O₃ (medium-term)?

Methods

Study design

Study sites were selected according to 1991-1993 annual mean O_3 levels in order to represent a broad spectrum of ambient O_3 exposure. To avoid confounding low levels of other pollutants from traffic and industry were required. Data was provided by governmental air pollution surveillance in both countries.

Initially 2,251 children were enrolled (Austria: 1150; Germany: 1101). Over a time period of 3.5 yrs two semi-annual lung function measurements were obtained (in spring and autumn) from schoolchildren initially in first and second primary school classes in Austria (Amstetten: altitude 270 m, St Valentin: 242 m, Krems: 190 m, Heidenreichstein: 560 m, Gänserndorf: 161 m, Mistelbach: 250 m, Wiesmath: 738 m, Bruck: 210 m, Pöllau: station site Masenberg at 1180 m) and South Western Germany (Aalen 420 m, Ehingen 530 m, Tuttlingen 643 m, Villingen 705 m, Freudenstadt 750 m, Welzheim 500 m). One German and 7 Austrian sites were rural and small town communities with a population size <12,000; 4 German and 2 Austrian sites were towns with a population size of 12,000–35,000; one German site had a population size of ~65,000.

The study was carried out from April 1994 to October 1997 in Austria and from February 1996 to October 1999 in Germany. Questionnaires, lung function and skin-prick test perfomance were standardised before the beginning of the study. In Germany, three teams of fieldworkers collected data. All lung function tests in Austria were performed by two trained medical students over the whole 3.5 yrs. Skin-prick tests were performed by a team of five doctors and medical students.

Teams and devices were changed systematically between study sites to avoid a bias caused by a fieldworker or device effect. All technical details (spirometry protocol, hardware, software, calibration of the spirometer) were standardised before the beginning of the study, and all staff were trained centrally. The study protocol was approved by the local ethics committee and written consent was obtained from the parents.

Pulmonary function testing

Pulmonary function tests took place during or after school time using a Masterscope (Software Release 4.0; Erich Jaeger, Würzburg, Germany). The instruments were calibrated daily. Every child performed two forced expiratory manoeuvres according to American Thoracic Society (ATS) guidelines, in a standing position wearing a noseclip. Two flow/volume curves were accepted as reproducible if the difference between FVC measurements was ≤5%. The highest FVC and FEV1 values were then selected for statistical analysis. Absolute values as well as values as per cent predicted (% pred) were reported (DOCKERY *et al.* [22]).

Questionnaires

Information concerning housing conditions, parental education, diagnosis of asthma or recurrent wheezy bronchitis

and history of respiratory symptoms was collected using questionnaires which were based on the standardised questions of the International Study of Asthma and Allergies in Childhood (ISAAC) questionnaire and filled in by the parents at the beginning of the study [23]. Further questionnaires concerning occurrence of respiratory symptoms during the study period were distributed each year and at the end of the study.

Skin-prick tests

Skin-prick tests were performed on the forearm with extracts from hazel, birch and grass pollens, cat and dog dander, house-dust mite, negative control: sodium chloride (9 g·L⁻¹), positive control: histamine hydrochloride (10 mg·mL⁻¹) (Alk Scherax, Hamburg, Germany). Skin reactions were evaluated after 15 min. A weal diameter >2 mm and at least half as large as the histamine weal was considered as a valid positive reaction [24].

Air pollution and exposure data

The concentrations of O_3 , nitrogen dioxide (NO_2) and total suspended particles (TSP)/particles with a 50% cut-off aerodynamic diameter of 10 μ m (PM10) were obtained as 30-min means from fixed monitoring stations located in the vicinity of the study site and run by local authorities. For O_3 the ultraviolet (UV) absorption method was used (Germany: KS-212-1012.10, Vorabscheider, Kalman System, Hungary; Austria: ML 8810 Monitor Labs, San Diego, CA, USA).

Statistical analysis

As a measure of lung growth [18] the difference between two consecutive FVC and FEV1 values for each child, divided by the number of days between the tests was calculated. The daily increase was then standardised for a period of 100 days (mL·100 days⁻¹). The effect of medium-term O₃ exposure on lung growth was analysed by means of linear regression models, with lung growth as the dependent variable and O₃ concentration averaged over each of the four summer periods (referred to as semi-annual mean O₃ exposure) as independent variable, further adjusting for sex, age and height at start of the time period, passive smoke exposure, short-term O₃ exposure at start and end of the time period, and time period. Short-term O₃ exposure was measured as maximum exposition on the day before lung function tests were performed. "Time period" consists of dummy variables for each season, thus reflecting effects that arise due to the longitudinal character of the study.

A regression model was constructed which described the relationship between O_3 exposure and lung growth directly, providing semi-annual O_3 means as independent variable. Semi-annual O_3 means differ between the 15 study sites and the four summer periods, but are equal within a given study site and study period. Thus the difference in lung growth caused hypothetically by a certain amount of increase in semi-annual O_3 concentrations was estimated.

Linearity of the relationship was investigated by means of the method of fractional polynomials [25]. Here, a set of polynomial transformations of the independent variable (semi-annual mean O₃ concentration) is provided, and the transformation fitting best is selected. Based on the results of the fractional polynomial selection precedure, exposure classes were formed with equal ranges of concentrations and similar numbers of observations. O₃ exposure classes were

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constructed separately for summer (April 1st until September 30th) and winter (October 1st until March 31st).

Further, an area-based categorisation with regions of high (H; 32.0–40.6 parts per billion (ppb)), medium (M; 24.1–30.7 ppb) or low (L; 20.0–21.3 ppb) exposure was used. The definition of these three O_3 exposure categories was arbitrarily made on the basis of O_3 means obtained for the total study period. In the regression model the area-based O_3 exposure categories served as independent and semi-annual lung growth as dependent variable. The intention of this analysis is to investigate sensitivity of the main model.

Further variables to be discussed as possibly influencing were medium-term NO_2 and sulphur dioxide (SO_2) exposition (mean values for the respective time period), fieldworkers performing the tests, and rhinitis observed when performing the test. Furthermore, a possible effect due to merging two studies and possible interactions of the "study effect" with an O_3 effect were taken into consideration.

Regression models were calculated first for each season separately. Then one common model was built for the four summer half-years and one model for the three winter half-years. This was done by applying the generalised estimation equation (GEE) method to account for the correlation of repeated measurements from each child [26]. To apply GEE, a so-called working correlation matrix must be specified. Estimates with asymptotically correct standard errors are then obtained independent of the choice of the working correlation. An exchangeable correlation structure was chosen which corresponds to a model including a random intercept term for each child. However, further choices (AR(1) and unstructured) were investigated.

Lung growth over the total study period of 3.5 yrs was assessed in a manner similar to that proposed by GAUDER-MAN and coworkers [15, 16]: in a first step, growth rates were obtained by performing individual regressions for log(FVC) and log(FEV1) as dependent variables, log(height) and short-term O₃ exposure as independent variables. This choice of modelling pre-adolescent lung growth has been strongly recommended by WARE and WEISS [27]. In a second step, individual regression slopes were analysed as dependent variables in a regression model with O₃ exposure averaged over four summer periods, sex, passive smoke exposure, and age at first lung function test as independent variables.

Only children with at least six lung function tests were included in these calculations in order to obtain reliable estimates covering a great part of the total observation period.

Results

With regard to medium-term effects, 2,153 children providing at least two consecutive lung function tests were analysed. There were in total 12,919 observations ("pairs" of two consecutive spirometries) on lung growth to be analysed. Of these 1,282 children had seven, 322 had six, 207 had five, 121 had four, 99 had three, 75 had two, and 47 had one pair of observations.

The main characteristics of the study population are described in table 1. Median age is 7.6 yrs on March 1st of the first study year; it is highest in Villingen and Aalen due to a larger proportion of children recruited from class two and lowest in Pöllau where investigations started later.

Asthma at the start of the study (defined as either diagnosed by a doctor or a positive answer to the ISAAC question "has your child ever had asthma") ranged between 1.2% (Heidenreichstein/Austria) and 12.8% (Amstetten/Austria). A possible explanation of such a broad range might be differing practices of local physicians when using the expressions "asthma" and "recurrent wheezing bronchitis". Sensitisation to pollen allergens also showed a large variation, between 6.8% (Wiesmath/Austria, Gänserndorf/Austria) and 28.8% (Ehingen/Germany).

In the total population passive smoke exposure was reported by 34.9% and ranged between 22.4% (Wiesmath/ Austria) and 46.5% (Krems/Austria). FVC % pred at the first investigation was 107.8% in the total population, ranging between 103.5% (Mistelbach/Austria) and 111.2% (Freudenstadt/Germany). FEV1 % pred ranged between 105.6% (St. Valentin/Austria) and 110.3% (Pöllau/Austria). Median FVC % pred (FEV1 % pred) was 110.3% (108.5%) in German regions and 105.0% (107.8%) in Austrian regions. Spearman correlation between FVC % pred and FEV1 % pred at the first investigation showed a similar relationship in Austria (r=0.82, p<0.0001) and in Germany (r=0.76, p<0.0001).

The allocation of the study sites into area-based categories (high (H), medium (M), low (L)) according to the mean O₃

Table 1. – Main characteristics of the study population

	Subjects n	Sex female %	Age years	Passive smoke exposure %	Asthma at start of the study %		FVC % pred [#] at start of study	FEV1 % pred [#] at start of study
Total population	2153	48.9	7.6 (6.5–8.8)	34.9	5.3	14.9	107.8 (89.6–125.6)	107.9 (88.4–126.6)
Amstetten (A)	134	43.3	7.5 (6.6–8.5)	40.5	12.8	13.3	105.9 (87.4–122.8)	107.7 (89.0–125.5)
St. Valentin (A)	116	51.7	7.7 (6.7–8.9)	31.3	8.8	12.6	104.3 (84.8–119.0)	105.6 (89.1–122.4)
Krems (A)	161	47.2	7.6 (6.6–8.5)	46.5	6.3	9.0	106.1 (89.5–127.7)	107.2 (87.0–126.0)
Tuttlingen (G)	231	48.1	7.8 (6.8–9.0)	45.8	3.5	20.2	107.8 (84.1–125.8)	107.1 (83.7–124.8)
Aalen (G)	89	53.9	8.3 (7.6–9.3)	42.7	4.8	17.9	110.3 (94.5–133.6)	107.6 (85.3–127.2)
Villingen (G)	35	37.1	8.3 (7.5–10.0)	38.2	2.9	11.8	107.7 (96.8–124.1)	109.4 (91.9–132.7)
Gänserndorf (A)	98	49.0	7.6 (6.7–8.6)	38.3	5.2	6.8	104.1 (83.7–131.1)	107.2 (91.2–123.6)
Ehingen (G)	73	54.8	7.7 (6.9–9.0)	32.9	2.7	28.8	111.0 (93.4–131.5)	108.5 (92.1–126.0)
Mistelbach (A)	152	44.7	7.6 (6.6–8.7)	30.3	4.0	8.9	103.5 (86.1–128.1)	105.4 (85.9–132.5)
Heidenreichstein (A)	83	50.6	7.6 (6.7–8.8)	32.9	1.2	11.6	105.5 (90.0–123.1)	107.7 (89.3–125.5)
Bruck (A)	73	52.1	7.5 (6.6–8.8)	43.1	4.1	14.1	103.9 (93.8–119.6)	104.5 (90.9–119.8)
Welzheim (G)	216	51.4	7.8 (6.8–9.0)	31.9	6.5	20.9	110.0 (95.8–125.3)	107.9 (88.9–128.6)
Freudenstadt (G)	395	49.4	7.7 (6.8–8.8)	28.7	3.3	18.1	111.2 (91.9–125.6)	109.4 (89.6–128.2)
Wiesmath (A)	126	47.6	7.6 (6.5–8.4)	22.4	4.0	6.8	103.8 (90.1–119.3)	109.1 (91.9–126.3)
Pöllau ⁺ (A)	171	49.7	6.6 (5.6–7.7)	30.6	8.2	8.1	106.5 (88.8–127.6)§	110.3 (87.1–129.9) [§]

Data are presented as median (5%–95% quantile) unless otherwise indicated. FVC: forced vital capacity; FEV1: forced expiratory volume in one second; A: Austria, G: Germany; +: Investigation at Pöllau started later (autumn 1994); \$\frac{\seta}{2}\$: data from 2nd investigation autumn 1994; #: predicted value calculated according to Dockery.

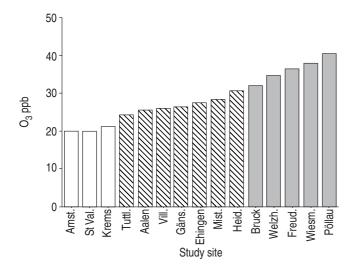


Fig. 1.—Mean 3.5 yearly ozone (O_3) concentrations (parts per billion (ppb)) in study sites divided into O_3 exposure categories: high $(H; \square)$, medium $(M; \square)$ or low $(L; \square)$ exposure.

concentration during the entire study period is shown in figure 1. Air pollution data stratified for the consecutive half years of the study period are presented in table 2. The O₃ concentrations of the half years (table 2) represent the same order as the data for the entire study period (fig. 1). Slightly higher O₃ values can be observed in the first two study summers (O₃ concentration averaged over all study sites is 36.6 ppb) than in the last two study summers (mean O₃ concentration is 35.1 ppb). Different measurements of particulate matters were employed in Austria (TSP) and Germany (PM10), and no measurements exist for Welzheim, Gänserndorf and Wiesmath (table 2).

Figures 2 and 3 show lung volume increases calculated as described above in mL·100 days⁻¹ for the area-based exposure categories: L, M or H exposure and for seven investigated study periods (four summer and three winter periods). In the

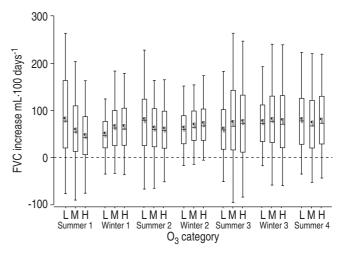


Fig. 2.—Increase in forced vital capacity (FVC) values over the seven study periods in the three area-based ozone (O_3) exposure categories (L: low; M: medium; H: high). The boxes include the interval between the 25%- and 75%-quantile, the extremes represent the 5%- and 95%-quantile, the median is shown as a horizontal line in the middle of the box and the mean as a symbol (\P).

low exposure category, a higher FVC increase in the first two summer periods and a smaller increase in the first two winter periods compared to the other exposure categories was observed (fig. 2). In the third summer period a lower increase in the low exposure category was found, and in the last two periods (winter three and summer four) no clear tendency for any of the exposure categories was detectable. Increases in FEV1 show a similar pattern over time (fig. 3).

In order to analyse the relationship between O_3 exposure and lung growth, multivariate regression models for the influence of O_3 exposure (area-based categories) on the lung volume increases were constructed for each time period separately (data not shown). The results of these regressions reflect those presented in figures 2 and 3.

Table 2. – Air pollution assessment: median of half hourly concentrations

	Amst.	St.Val.	Krems	Tuttl.	Aalen	Vill.	Gäns.	Ehing.	Mistel.	Heid.	Bruck	Welz.	Freud.	Wies.	Pöllau
O ₃ summer1	24	27	30	30	27	37	35	31	35	38	41	41	44	49	44
O ₃ winter 1	5	10	12	6	9	6	18	12	18	23	22	22	24	32	31
O ₃ summer 2	25	23	24	31	34	31	31	36	35	35	36	44	43	46	48
O ₃ winter 2	9	10	9	12	13	15	15	14	19	23	19	22	29	25	34
O ₃ summer 3	22	21	21	35	36	34	29	36	31	34	36	41	44	41	45
O ₃ winter 3	7	8	8	14	13	17	16	18	16	21	23	23	31	25	32
O ₃ summer 4	27	17	17	30	29	36	31	33	33	37	38	44	43	39	49
NO ₂ summer1	11	8	11	9	9	7	7	7	6	5	4	3	5	3	2
NO ₂ winter 1	17	10	15	18	17	13	10	14	8	5	7	9	9	8	3
NO ₂ summer 2	10	9	11	9	11	7	5	7	5	3	4	4	6	1	3
NO ₂ winter 2	16	14	16	17	18	14	10	13	8	6	8	8	8	3	4
NO ₂ summer 3	10	5	4	8	10	7	5	7	6	4	3	4	5	5	2
NO ₂ winter 3	10	7	12	17	16	13	10	13		5	7	9	8		2
NO ₂ summer 4	7	4	2	8	10	6	6	7		4	4	4	6		2
PM10/TSP sum.1	30	26	30	14	19	16		19			30		15		15
PM10/TSP win.1	20	26	30	18	30	17		21			20		14		7
PM10/TSP sum.2	20	25	30	14	26	15		17	20		20		16		15
PM ₁₀ /TSP win.2	40	35	40	17	27	19		21	35		50		15		11
PM ₁₀ /TSP sum.3	20	24	30	13	22	17		15	20	20	20		16		14
PM ₁₀ /TSP win.3	40	30	30	17	24	24		17	20	20	30		12		9
PM10/TSP sum.4	30	23	30	12	21	17		15	20	20	20		15		14

Ozone (O₃) in parts per billion (ppb), nitrogen dioxide (NO₂) in ppb; total suspended particles (TSP) and particles with a 50% cut-off aerodynamic diameter of 10 μm (PM10) in μg·m⁻³; Sum: summer; win: winter; TSP data available for Amstetten, St.Valentin, Krems, Mistelbach, Heidenreichstein, Bruck, Pöllau; PM10 data available for Tuttlingen, Aalen, Villingen, Ehingen, Freudenstadt.

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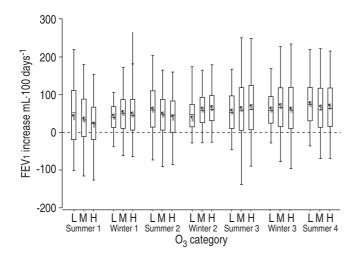


Fig. 3.—Increase in forced expiratory volume in one second (FEV1) increase values over the seven study periods in the three area-based ozone (O₃) exposure categories (L: low; M: medium; H: high). The boxes include the interval between the 25%- and 75%-quantile, the extremes represent the 5%- and 95%-quantile, the median is shown as a horizontal line in the middle of the box and the mean as a symbol (¶).

Table 3 shows the common regression analysis of the four summer periods and the three winter periods, analysing lung volume increases as the dependent variable and semi-annual mean O_3 concentrations as the independent variable of interest. Models were adjusted for sex, age and height at start of the time period, passive smoke exposure, short term O_3 exposure and time period as further independent variables.

Linearity of the relationship between semi-annual mean O₃ concentrations and lung growth was investigated *via* fractional polynomials with the same type of regression model as above, but systematically introducing nonlinear transformations

Table 3. – Parameter estimates from generalised estimation equation models* for semi-annual ozone (O₃) means, divided in classes (pooled analysis for seasons)

Increase in:

Parameter estimates

for increase (mL·100 days⁻¹)

	for semi-annual O ₃ means compared to "46–54 ppb" (summer) or to "28–36 ppb' (winter)
FVC summer	
22–<30 ppb	19.2 (10.6–27.8)
30–<38 ppb	15.5 (7.0–24.0)
38–<46 ppb	19.2 (11.1–27.3)
FEV1 summer	
22–<30 ppb	18.5 (9.8–27.1)
30–<38 ppb	11.6 (3.2–20.1)
38–<46 ppb	16.8 (8.5–25.0)
FVC winter	
4-<12 ppb	-16.4 (-24.68.3)
12-<20 ppb	-15.0 (-21.68.4)
20–<28 ppb	-11.8 (-18.45.3)
FEV1 winter	
4–<12 ppb	-10.9 (-19.72.1)
12-<20 ppb	-11.2 (-18.53.9)
20–<28 ppb	-9.9 (-17.42.4)
* *	` /

Data are presented as parameter estimate (95% confidence intervals). FVC: forced vital capacity; FEV1: forced expiratory volume in one second; ppb: parts per billion. $^{\#}$: Adjusted for sex, age and height at start of the time period, passive smoke exposure, time period, short-term O_3 exposure at start and end of time period.

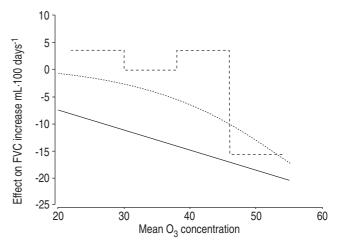


Fig. 4.—Estimated functional influence of mean ozone (O₃) exposure on forced vital capacity (FVC) increase in summer: —: linear; ---: classified; ·----: fractional polynomials.

of semi-annual mean O_3 concentration. The estimated relationship between semi-annual mean O_3 concentration and FVC increase in summer is shown in figure 4 together with an estimated linear relationship and a division of semi-annual mean O_3 concentrations into four classes.

Based on these results, analysis was performed by dividing semi-annual mean O₃ concentrations into four classes and revealed significantly higher increases in FVC and FEV1 during summer for all classes in comparison to the highest exposure (46–54 ppb) (table 3). For example, the semi-annual mean O₃ exposure in summer of 22–30 ppb (compared to 46–54 ppb) was associated with an estimated higher increase in FVC (FEV1) of 19.2 (18.5) mL·100 days⁻¹ (table 3). An inverse effect occurred for the winter periods. For instance, the semi-annual mean O₃ exposure in winter of 4–12 ppb (compared to 28–36 ppb) was associated with an estimated lower increase in FVC (FEV1) of -16.4 (-10.9) mL·100 days⁻¹ (table 3).

Interaction between a possible "country effect" and the possible O₃ effect was tested and found not to be present. Introducing medium-term NO₂, medium-term SO₂, rhinitis when performing the test, or fieldworkers as a possibly influencing variable did not change the results concerning O₃ effects and were therefore not considered in the final models.

To identify combined effects, stratified analyses of the final O₃ exposure class model were conducted, stratifying the population with respect to sex, pollen sensitisation, asthma, and a diagnosis of hayfever. No significant interaction could be found (data not shown).

An age-stratified analysis was performed to investigate the possibility of a higher vulnerability in younger children, where groups of approximately equal size were formed, separated according to their age at the first investigation. No differences could be found for summer effects, but in the analysis of winter periods the effect was significantly less pronounced in younger children (data not shown).

For comparative reasons, the analysis was repeated using area-based categories, where a significantly higher increase during summer was found in the low but not in the medium O_3 exposure category compared to the high exposure category for both FVC (8.0 mL·100 days⁻¹) and FEV1 (9.3 mL·100 days⁻¹; table 4). In contrast, increases in winter were significantly lower in the low exposure category compared to the high exposure category (-9.7 mL·100 days⁻¹ for FVC and -7.5 mL·100 days⁻¹ for FEV1).

To investigate the hypothesis that a positive O_3 effect in winter can be viewed as a catch-up effect of a preceding negative O₃ effect in summer, regression coefficients were calculated where summer FVC/FEV1 increases are used to predict increases in the succesive winter. The regression coefficients were estimated separately for O₃ exposure areas (L/M/H), and the parameter estimates (PE) and their standard errors (SE) are for FVC increase: PE=-0.165, SE=0.022 (L); PE=-0.154, SE=0.025 (M); PE=-0.319, SE=0.031 (H) for summer 1, winter 1; PE=-0.240, SE=0.026 (L); PE=-0.283, SE=0.026 (M); PE=-0.344, SE=0.025 (H) for summer 2, winter 2; PE=-0.351, SE=0.046 (L); PE=-0.319, SE=0.029 (M); PE=-0.344, SE=0.030 (H) for summer 3, winter 3. The PE for the summer 1/winter 1 model is markedly higher in the high exposure area, whereas PEs are of similar size for the following seasons. This effect is even more prominent when analysing FEV1 increases (data not shown).

For 1,869 children with at least six lung function tests, individual regression coefficients for log(FVC) and log(FEV1) regressed on log(height) and short-term O₃ exposure were obtained. The individual regression slopes were then related to mean summer O₃ exposure over four periods, sex, passive smoke exposure, and age at first lung function test as independent variables in a second regression model. No association between growth rates and mean summer O₃ could be shown for FVC and FEV1 over a 3.5 yearly period.

Discussion

Based on 12,919 observations of schoolchildren's lung growth, higher increases in FVC and FEV1 were observed in the case of low ambient O₃ exposure during summer when high exposure served as a reference. In turn, during winter, exposure to lower ambient O₃ concentrations was associated with a lower increase in FVC and FEV1 in comparison to exposure to high concentrations. However, low exposure during summer represents low exposure also during winter. Thus the current authors postulate that the winter results are merely a reversibility pattern regarding O₃ effects during summer. This is especially the case for the area-based analysis (table 4), where almost the same individuals were in the same exposure category for both summer and winter. That would mean that only the category with high exposure during summer (H; 3.5-yr average 32.0-40.6 ppb) differs from the other groups, and during winter decreasing exposure reveals opposite differences in the changes of FVC and FEV1. Results from separate seasonal analyses show however no consistent effects: there was a clear effect observed in the first

Table 4. – Parameter estimates from generalised estimation equation models[#] for ozone (O₃) exposure categories (pooled analysis for seasons) with "high exposure" as reference

Increase in:

Parameter estimates for increase
(mL·100 days⁻¹) for O₃ exposure categories
compared to "high"

	Medium	Low
FVC summer	-0.03 (-5.0-4.9)	8.0 (1.8–14.3)
FVC winter	-5.8 (-10.31.3)	-9.7 (-14.5–4.9)
FEV1 summer	-0.1 (-5.35.1)	9.3 (3.0–15.6)
FEV1 winter	-3.1 (-8.2-2.0)	-7.5 (-12.7–2.4)

Data are presented as parameter estimate (95% confidence intervals). FVC: forced vital capacity; FEV1: forced expiratory volume in one second. $^{\#}$: Adjusted for sex, age and height at start of the time period, passive smoke exposure, time period, short-term O_3 exposure at start and end of time period.

two study years and a much broader variability of effects in the later part of the study period. These findings point to a possible medium-term effect (in summer seasons) of high ambient O_3 concentrations on schoolchildren's lung growth, which seems to be transitoric.

The current authors have presented a pooled analysis of two data sets collected in Austria (1994-1996) and Germany (1996–1999). FVC % pred was found to be higher in German regions but no clear differences for FEV1 % pred could be observed. The pooling of different data sets might thus give rise to study effects in the analysis. However, all technical details (spirometry protocol, hardware, software, calibration of the spirometer) were standardised before the beginning of the study, and all staff were trained centrally. The fact that differences in lung function measurements could only be observed for FVC % pred but not for FEV1 % pred further hints at the absence of systematic differences. Interaction terms of the factors "study" and "O3 exposure" were considered in the statistical analysis and were found to be insignificant for the main results (data not shown). The pooled analysis of two large data sets thus has the advantage of greatly increasing the power of the analysis. Furthermore, it was possible to include a broader range of ambient O₃ exposure in the analysis, leading to more stability in the statistical parameter estimates.

In earlier studies assessment of medium-term (referring to months) or long-term effects (referring to years) of air pollution was often done by conducting a cross-sectional study where people from differently exposed areas were compared [28, 29]. The longitudinal design of this study offers the opportunity of comparing lung growth measurements of the same child to each other, which reduces in particular the possibility of a migration bias (people who become ill will tend to migrate to healthier areas [30]).

One major problem in the assessment of adverse health effects of air pollutants is the choice of an adequate exposure variable. In order to address this problem, two approaches were chosen, one consisting of dividing the study regions into three exposure categories low/medium/high, and one where semi-annual mean O₃ concentrations were used. Defining high or low exposed geographical areas as a surrogate variable for air pollution exposure is a robust but rough means: it might be the source of potential bias caused by the confounding of other air pollutants. In order to obtain an exposure variable that is more specific, mean semi-annual O₃ concentrations were calculated. Especially in the context of a longitudinal study this approach allows exposure which varies from season to season to be taken into consideration.

Stationary monitors were used to assess air pollution exposure. Some authors have described dependencies between air pollution effects and time spent outdoors [15, 16]. More accurate measurements could possibly have been obtained by the use of personal samplers which seems however hardly feasible for a large study population over such a long study period. The present authors have investigated the association between stationary monitor measurements, personal sampling, and short-term effects of these measurements on lung function in a subgroup of the study population [31] and hence consider that stationary monitors provide valid measurements.

The shape of the function chosen to describe the O_3 exposure effect was considered carefully. Transformations of the independent exposure variable in the linear regression model were investigated systematically and revealed a nonlinear relationship between O_3 exposure and lung growth. In general, the most robust method is the classification of semi-annual mean O_3 concentrations into several classes and was applied to the current study's data. The comparison of the two O_3 exposure assessments (considering area-based categories *versus* semi-annual O_3 concentrations) shows that

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the same conclusion is reached, namely transitoric effects of O_3 exposure on lung function, but the details of the evaluation appear different: table 3 shows that in the first model (O_3 concentrations) effects are outstanding for all "lower" classes (22–46 ppb) in comparison to the highest class (46–54 ppb; table 3); table 4 shows that for the second model (area-based) effects are outstanding for the low exposure category in comparison to the reference with high exposure. Thus in general there is no contradiction. The approach using O_3 concentrations underlines a limit value of approximately 46 ppb.

The effect of other air pollutants such as NO₂, SO₂, particulate matter (PM10) or TSP on children's lung growth should also be investigated as it is a possible source of biased estimation. Therefore the final analysis was repeated (table 3) by introducing mean SO₂ and mean NO₂ values as confounding factors. The parameter estimates for O₃ were very similar to the results presented in table 3. It was concluded that the effect of other air pollutants on children's lung growth may be present but does not interfere with the effects describing at present. Concerning PM10/TSP, no measurements were taken in three of the study sites. The present authors therefore refer to earlier work [18, 21, 32] where effects of PM10 exposure did not alter the findings with respect to O₃.

Short-term effects of exposure to O₃ have been extensively studied [1–13], whereas only few investigations dealing with medium-term or long-term effects have been published [14–18, 21].

Long-term effects of O_3 on lung growth have been assessed in animal studies, where subchronic effects of O₃ were observed in rats using markers of cell injury [33–36]. TYLER et al. [33] concluded that O₃ inhalation by young rats altered lung growth and development in ways likely to be detrimental, and that those changes persisted after O₃ exposure had been terminated [34]. VAN BREE et al. [36] found that some of the markers (protein and albumin content, neutrophil influx in bronchoalveolar fluid) returned to control levels within a few days after termination of exposure, whereas other markers (alveolar macrophage response, presence of terminal bronchioles, thickening of ductular septa by enhanced cellularity, collagen formation) were persistently elevated even after exposure. Thus, inflammatory processes could result also in long-term effects on lung growth even under changing exposure. Assuming a cumulative effect of natural exposure, ACKERMANN-LIEBRICH et al. [28] found no signs of O₃ influences on FVC and FEV1 in adulthood. In contrast, the cross-sectional analysis of a study in South California showed a statistically significant association between O₃ exposure and lower FVC and FEV1 for male children spending much time outdoors and for female children with asthma [14]. Based on a longitudinal design KINNEY et al. [17] observed a larger decline in FEV1 (44 mL) in the area with highest O₃ exposure over one summer period in healthy young adults working outdoors. The estimated average O₃ effects of the Austrian study by Frischer et al. [18], which constitutes a part of the current authors' data, were -1.8 (-2.9) mL·100 days⁻¹·ppb⁻¹ increase in FVC (FEV1) in three summer periods. The average summer O₃ effect observed by the current authors corresponds to an effect of about -0.80 (-0.78) mL·100 days⁻¹·ppb⁻¹ increase in FVC (FEV1) if it were considered to be linear, or -19.2 (-18.5) mL·100 days⁻¹ for high exposure (46-54 ppb) compared to low exposure (22–30 ppb). They are thus slightly weaker than the effects described by Frischer et al. [18]. In an earlier article the current authors' group reported summer O₃ effects on the difference in FVC (FEV1) increase of -27.8 and -26.6 (-30.3 and -32.2) mL·100 days⁻¹, when comparing high exposure areas to low exposure areas over two summers

[21]. Thus the current results of negative effects of O_3 averaged over four summer seasons are in accordance with earlier results and speak in favour of a medium-term effect on schoolchildren's lung growth.

However, the O_3 effects observed during winter periods are of roughly the same size and show reversed signs: a lower increase in FVC (FEV1) of -16.4 (-10.9) mL·100 days⁻¹ for low exposure (4–12 ppb) compared to high exposure (28–36 ppb). Regression coefficients where summer FVC/FEV1 increases are used to predict increases in the successive winter were calculated to support the current authors' view that the winter O_3 effect can be interpreted as a catch-up effect. The estimated regression slope was found to be markedly higher in the high exposure area than in the other areas for the first season (summer 1, winter 1), but not for the following seasons. These regression models were used to relate summer and winter FVC (FEV1) increases of the children directly and indicate that winter O_3 effects can be viewed as a reversibility effect.

In addition these results are in accordance with those of FRISCHER *et al.* [18] and KOPP *et al.* [21]. Longer-lasting effects of O_3 on children's lung growth thus seem to be at least questionable. The longitudinal analysis of a South Californian cohort [15] points in the same direction, showing no signs of an O_3 effect over a time period of 4 yrs, and although the results of the second cohort [16] showed effects of O_3 on changes in PEFR (difference in annual per cent growth rates from the least to the most polluted community is -1.21%) it found none on changes in FVC and FEV1.

There were varying results from the seasonal analyses. A clear effect of O₃ on children's lung growth in the first 2 yrs of the study was not repeated in the later part of the study. Two possible explanations can be taken into consideration. One reason might be that younger children are possibly more vulnerable than older children, an assumption which can also be supported by the findings of the current authors' earlier studies [18, 21], where consistent effects were found for the first 2-yr study period but not in the analysis of the third study year of the Austrian data. The age-stratified analysis revealed a significantly less pronounciated effect in winter for younger children. Since "winter ozone effects" are interpreted as a catch-up of harmful summer effects, this result can be viewed in support of the hypothesis that younger children are more vulnerable.

Another reason might be the variation of O_3 exposure during different study summer periods, where especially in the Austrian high O_3 exposure regions the highest mean O_3 concentrations were observed during the first part of the study. This aspect again emphasises the importance of accurate exposure measurement. The common analysis of the whole study period in this case presents an estimated averaged O_3 effect over summer periods with different exposure levels.

Conclusion

An effect of summer ozone exposure on the growth-related increase in forced vital capacity and forced expiratory volume in one second in schoolchildren was observed only for the highest ozone concentrations during the first 2 yrs of the study. In the following winter seasons the pattern was reversed. It is thus concluded that medium-term effects on schoolchildren's lung growth are possibly present. However, in the long-term no association between growth rates and mean summer O₃ could be shown for forced vital capacity and forced expiratory volume in one second. Medium term effects thus might be at least partially reversible.

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